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## **Impressum:**

CESifo Working Papers

ISSN 2364-1428 (electronic version)

Publisher and distributor: Munich Society for the Promotion of Economic Research - CESifo GmbH

The international platform of Ludwigs-Maximilians University's Center for Economic Studies and the ifo Institute

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# The Effect of Language Training on Immigrants' Economic Integration Empirical Evidence from France

## Abstract

We examine the impact of language training on the economic integration of immigrants in France. The assignment to this training, offered by the French Ministry of the Interior, depends mainly on a precise rule: the training is available when the test score of an initial language exam is below a certain threshold. This eligibility rule creates a discontinuity in the relation between the test result and the variables of interest, which is used to estimate the causal effect of this training, through the method of Regression Discontinuity Design. We find that the number of assigned hours of training significantly increases labor force participation of the treated individuals. The language classes appear to have a larger effect for labor migrants and refugees relative to family migrants, for men and individuals below the median age, and for individuals with higher levels of education. Our estimates suggest that the main channel for the improved labor market participation is the information on job search strategies that immigrants derive from the interaction with their classmates and teachers during classes.

JEL-Codes: J150, J610, J680.

Keywords: immigrants' integration, language training, Regression Discontinuity Design.

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April 23, 2017

This research project benefitted from the support of the Département des statistiques, des études et de la documentation (DSED) at the Direction générale des étrangers en France (DGEF) of the French Ministry of the Interior. We are particularly indebted to Marie-Hélène Amiel and Gérard Bouvier, DSED's director and deputy-director. This research was also supported by the "Flash Asile" program of the Agence Nationale de la Recherche (ANR-16-FASI-0001). We thank Toman Barsbai, Béatrice Boulu, Gabrielle Fack, Benjamin Feigenberg, Jesús Fernández-Huertas Moraga, Delia Furtado, Christina Gathmann, Jeffrey Grogger, Ricardo Hausmann, Virginie Jourdan, Ole Monscheuer, Ljubica Nedelkoska, Frank Neffke, Francesc Ortega, Xavier Pauly, Paolo Pinotti, Marco Ranaldi, Melissa Siegel, and Andreas Steinmayr for their helpful comments and suggestions. We are also grateful to conference and seminar audiences at the 2017 Royal Economic Society Annual Conference in Bristol, CEPII-OECD 6th Annual International Conference, University of Mannheim, RWI Essen, UNU-MERIT and Center for International Development (CID) at Harvard University.

# 1 Introduction

Immigrants often represent the most vulnerable part of the labor force. Amongst others, Algan et al. (2010) show that France does not represent an exception. Following a social and economic degradation of immigrants' situation and in the face of pervasive discriminatory practices regarding work and housing, the French Government has introduced policies aimed at improving the economic and social integration of immigrants and at providing immigrants the dignity guaranteed by fundamental human rights. Since the 1st of January 2007, every new legal immigrant to France who is older than 16 and is coming from a country outside the EU, has to sign a *Contrat d'accueil et d'intégration* (CAI). This contract imposes a civil training (e.g., on French institutions and the values of the French Republic), a language training, an information session on life in France and a statement of professional competence.<sup>1</sup>

This study evaluates the component of the CAI related to the language training. In fact, after the signature of the CAI contract, the immigrant has to pass a test on the knowledge of the French language, written and spoken. If the result is insufficient, the person is likely to receive a training. The assignment to this training depends, therefore, mainly on a precise rule: the treatment (i.e., the language training) is available when the test results are inferior or equal to a certain threshold. This eligibility rule creates a discontinuity in the relation between the test result and the variables of interest. We use this discontinuity to estimate the causal effect of the CAI language training, through the method of Regression Discontinuity (RD) design. Straightforward regression analysis could lead to bias in the estimation of the relation, since individuals who take the language classes are different in unobservable characteristics from the individuals who are not assigned the training. The proposed method is based on the hypothesis that immigrants who are just above and just below the eligibility threshold are comparable and similar in several dimensions, except for their participation in the linguistic training organized by the Ministry of the Interior. This unique framework allows, thus, to assess one major factor that impacts the immigrants' efficient integration in the host country's labor market: the language barrier. Chiswick (1991) and Borjas (1994) were among the first scholars to recognize the language barrier as one of the major hurdles in the immigrant's integration. Since then, improving the language skills has been acknowledged to play an important role in the assimilation

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<sup>1</sup>On the 1st of July 2016, the French Government modified some aspects of these integration policies. The new integration plan is denoted *Contrat d'intégration républicaine*, CIR.

and integration process. Indeed, the data we use in our analysis – the ELIPA dataset – show that 45% of all individuals in our sample think that the lack of French language skills is the main obstacle to immigrants' integration.

The econometric analysis of the relationship between language training and economic integration is challenging, because if the least skilled immigrants attend the language classes and some skill dimensions are unobservable, then traditional econometric tools would underestimate the true effect of the training. Our study is unique in its novel approach towards the assessment of language classes on several outcomes related to immigrants' integration. We use a local randomized experiment in the form of a Regression Discontinuity Design to correct for the selection-into-language-training due to unobservables, such as ability.

Our empirical analysis shows a higher probability of participating in the labor force due to the language classes. The mechanisms hint towards an information effect: before, during and after the classes, immigrants can obtain useful information on job search strategies from interactions with classmates and teachers. Looking at the heterogeneity of effects on labor force participation by education level, we find that the higher the level of education, the bigger the impact of the language training on labor force participation. Illiterate individuals have little or no benefit from the language classes. Furthermore, there appears to be a higher effect of the training on labor force participation for labor migrants and refugees, as compared to family migrants. Women tend to benefit less from the language classes in terms of labor force participation. The same is valid for individuals above the median age. Instead our estimates do not show significant improvements in objective measures of language skills – which might be due to the basic level of the CAI French classes – nor changes in the probability of employment for the immigrants who were assigned the language training.

Our paper is related to several branches of literature. There are works on the relationship between language skills and labor market performance, see among others Chiswick (1991), Chiswick and Miller (1995), Dustmann and van Soest (2001), Dustmann and Fabbri (2003), and Bleakley and Chin (2004). Chiswick (1991) studies the determinants of English language fluency among immigrants in the US and the effects of self-reported language skills on earnings. He finds that reading fluency is more important than speaking fluency as a determinant of income. Chiswick and Miller (1995) provide an international comparison of the relationship between language and income, using data from Australia, the US, Canada and Israel. Dustmann and van Soest (2001) show that neglecting measurement er-

ror in self-reported assessments of language proficiency leads to a substantial downward bias of the impact of speaking fluency on earnings. For nonparametric identification of the earnings equation, they use the education level of the immigrant's father as instrument in the speaking fluency equation. Dustmann and Fabbri (2003) analyze data from the UK. They combine a matching estimator that addresses the problem of endogenous choice of language acquisition with an IV estimator that eliminates the bias due to measurement error in the self-assessed measures of language skills. The instrument they use is a dummy variable equal to one if the interview was done in English only. Their estimates show that fluency in English increases employment probabilities by about 22 percentage points. Instead, the estimates on earnings are not statistically significant, when both endogenous selection and measurement error are taken into account. Finally, Bleakley and Chin (2004) solve the endogeneity of language skills in immigrants' earnings equations, using age at arrival interacted with a dummy for non-English-speaking country as instrumental variable. We complement this branch of the literature by exploring how integration plans proposed by the government affect labor market integration, not only through their potential effect on objective measures of language skills, but through other channels as well. We also innovate using a RD design, which has the properties of a local randomized experiment.

The second branch of the literature this paper is related to, uses test scores or other eligibility rules as assignment variables of local randomized experiments in contexts of human capital investment. Matsuura (2008) exploits the fact that students are assigned to a summer school based on their score on year-end exams and shows that summer schools may be a less costly alternative to class-size reductions when trying to increase student achievement. Urquiola and Verhoogen (2009) study the effect of class-size on student outcomes, focusing on the relationship between schools' choices of class size and households' choices of schools.<sup>2</sup> Manacorda (2012) employs data on junior high school students in Uruguay and finds that grade failure increases subsequent dropout and reduces subsequent educational attainment. Chin et al. (2013) investigate the effect of bilingual education on the academic achievement of students with low English proficiency and their peers in Texas. In case of bilingual education, the students attend some classes in English and other classes in the native language (with Spanish-English programs by far the most common). The authors exploit a policy rule requiring a school district to offer bilin-

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<sup>2</sup>Another work analyzing class size effects is Angrist et al. (2016).

gual education depending on a cutoff value of enrollment of limited English proficient students.<sup>3</sup> Our work contributes to this literature by focusing on the effects of language training for immigrants, which is a relevant research question given that poor language skills is a major obstacle for the immigrants' economic and social integration in host countries.

Finally, the paper that comes closer to our study is by Sarvimäki and Hämäläinen (2016). The authors analyze the effect of a change in Active Labor Market Policies for unemployed immigrants in Finland, using a fuzzy RD design that relies on a cutoff date of arrival determining whether immigrants are affected or not by the reform. Sarvimäki and Hämäläinen (2016) estimate the effect of implementing tutoring of caseworkers in integration plans, which aims to improve the communication between caseworkers and immigrants and to make sure the immigrants understand their integration plan. Over a follow-up period of 10 years, they find remarkable benefits of the tutoring policy for the immigrants. Differently from that study, our work focuses on the impact of the language training itself (rather than tutoring of caseworkers in integration plans), on a broader set of outcomes related to the immigrant's integration (among others, the formation of networks and objective measures of language skills), and on the test scores of an initial language exam as assignment variable (rather than a cutoff date determining eligibility for the reform). Moreover, we empirically assess several mechanisms that might explain the relationship between languages classes and immigrants' economic integration.

The rest of the paper is structured as follows. Section 2 provides background information on immigrants' integration policies in France. Section 3 presents the data this study relies on and descriptive statistics that motivate our analysis. Section 4 explains the empirical strategy. Section 5 presents the results, as well as the mechanisms at work. Finally, section 6 concludes.

## 2 The French "*Contrat d'Accueil et d'Intégration*" (CAI)

As previously explained, every new legal immigrant to France who is older than 16 and is coming from a country outside the EU, has to sign a contract, which imposes a civil

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<sup>3</sup>See Cappellari and Di Paolo (2015) for a study of bilingual schooling and earnings in Catalonia, which relies on a difference-in-differences framework and exploits a reform that introduced Catalan alongside Spanish as medium of instruction in Catalan schools. In addition to the articles summarized in this introduction, examples of other works that study human capital investment using a RD design are Van der Klaauw (2002) and Jacob and Lefgren (2004).

training (e.g., on French institutions and on the values of the French Republic), a language training, an information session on life in France and a statement of professional competence. Until recently this integration contract was the so-called *Contrat d'accueil et d'intégration* (CAI), introduced in 2007. In July 2016, the French Government launched an updated version of this program, under the name of *Contrat d'intégration republicaine* (CIR). In this paper, we evaluate the language training component of the CAI.

Following the signature of the CAI, all individuals had to take a language test. This language test did not evaluate immigrants' actual language skills, but their skills relative to an A1.1 (a "survival") level. Individuals with language skills above this level were likely to succeed at the highest grade, even though their French language skills might have been very basic. The number of assigned hours of training ranged from 60 to 400, depending on how much training was needed to achieve an A1.1 level. Once the training was over, an evaluation was made and the individuals got a diploma of basic French skills, which was an important requirement to renew the residency permit. After the first renewal, this one-time language course allowed to extend the residency permit as often as the individuals wished. Only in case the immigrants wanted to naturalize, they had to present a B1 certificate level. In sum, the CAI system foresaw no system of progressive learning of the new language: once the level A1.1 was attained, individuals did not have to further improve their French - unless they wanted to apply for naturalization. Thus the offered training stopped at levels that were generally too low for economic integration. The individuals did not speak French well enough to be considered employable.

While our work focuses on the evaluation of the language training of the CAI, it is interesting to compare the old version of the integration policies with the new version of the contract. The CIR too imposes the language entry test to all individuals who sign the contract. In contrast to the CAI, however the language test of the CIR evaluates the actual language skills, to position individuals in courses according to their skills at levels A1, A2 or B1. The number of assigned hours varies between 50, 100 and 200 hours - according to the needs of the individual to achieve one of these three levels. Once the training is over, an evaluation of at least the basic level A1 is carried out. In order to renew the residence permit, the immigrants must show progress in speaking and writing French. If the individuals progress, they can move from the 1-year-residence permit to the "multi-annual" permit (2 to 4 years, a new type of residence permit introduced with the CIR). As a next step, the individuals must follow a free A2 level course offered by the government. This time, to renew the "multiannual" residency permit, the individuals must successfully



complete an exam that shows their A2 French skills. Finally, only if individuals want to naturalize, they need to have a B1 level. The French government offers classes for this level as well. The major difference between the CAI and the CIR's language classes concerns the system of progressive learning of French in the latter training, which follows the individuals up to the B1 level and allows them to learn French at employability levels.

### 3 Data

The study is based on the database *Enquête Longitudinale sur l'Intégration des Primo-Arrivants* (ELIPA), created by the *Département des statistiques, des études et de la documentation* (DSED) of the French Ministry of the Interior. These data contain detailed socio-demographic information on the interviewed immigrants. For each individual in the sample, the dataset also includes administrative information by the *Office français de l'immigration et de l'intégration* (OFII) on the test results that determine admission to the language training program.

#### 3.1 ELIPA

The *Enquête Longitudinale sur l'Intégration des Primo-Arrivants* (ELIPA) is a longitudinal survey on the integration of first-time arrived immigrants in France. 6107 individuals were interviewed and followed over three periods of time. The first wave was carried out in 2010 (baseline survey), the second wave in 2011, and the final wave took place in 2013. The survey aims at following the path of the individuals and collects a large set of information regarding socio-demographic characteristics, bureaucratic itinerary, employment, language skills, living conditions and social integration. The individuals are immigrants from countries outside the EU and Switzerland who are at least 18 years old and were granted their residence permit at the end of 2009.<sup>4</sup> The survey considers individuals who asked for a residence permit of at least one year and want to settle permanently in the

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<sup>4</sup>The dataset only includes information on immigrants residing in the host country with a regular residence permit. Examples of works that analyze the effects of legal status on several aspects of immigrants' integration are Kossoudji and Cobb-Clark (2002) and Dustmann, Fasani and Speciale (forthcoming).

country.<sup>5</sup>

In order to evaluate the integration of immigrants as broadly as possible, we use several different outcome variables.

### 3.2 Data on Test Results and Outcome Variables

We measure labor market outcomes in 2013, three years after the baseline survey and the entry test determining assignment to the language classes. Our main dependent variables are labor force participation (equal to 1 if being employed or registered as unemployed, 0 otherwise)<sup>6</sup>, employment status (equal to 1 if employed, 0 otherwise), type of contract (equal to 1 if permanent contract) and type of employment (a dummy equal to 1 if full time job; another dummy equal to 1 if informal work). Furthermore we also measure income per household individual.<sup>7</sup>

The ELIPA dataset provides detailed information on the language training and results of the entry test. The latter information is not self-reported, but it is provided by the OFII. The training is not assigned to all individuals. It is exclusively offered to, and compulsory for, a subset of individuals. After signing the CAI, each immigrant has to take a language test, which is composed of five parts: one oral part that amounts to a total of 70 points maximum and four written exams that, together with the oral exam, amount to a total of 100 points. We create a running variable that sums the points gained from each of the tests. The total grade is in intervals of five points and its maximum value is 100 points. 45 is the last failing grade and 50 is the first passing grade.

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<sup>5</sup>The following categories of residence permits are excluded from the survey: students and trainees; employees on assignment; holders of residence permits with endorsements for "skills and talents", "research" or "business"; high-level executives and the family members accompanying them; holders of a permit endorsed as "visitor" or "arts and culture professional"; veterans; and persons from third countries entering France for medical treatments (Direction générale des étrangers en France, Département des statistiques, des études et de la documentation, 2010).

<sup>6</sup>In the section where we present the results, we mainly focus on labor force participation because this is the dependent variable for which we find statistically significant effects of the language classes.

<sup>7</sup>Income per household individual is defined as the total income of the household divided by its "consumption units". These are calculated following the equivalence scale of the OECD, which attributes one consumption unit to the head of the household, 0.5 to other members of age 14 and older, and 0.3 to children below the age of 14.

### 3.3 Descriptive Statistics

Most of the new immigrants (69%) were granted a residence permit on the grounds of family migration. The other two main administrative reasons are the granting of refugee status (14%) or economic migration, which is migration for employment reasons (12%). The slight majority of the new immigrants (51%) are women. The average age at arrival is 29. All the individuals in the sample received their residence permit in 2009, however, they have not all arrived in the same year: on average, immigrants have been living in France 3.5 years before the survey. The main region of origin is Africa: 33% are from the Maghreb and 28% from sub-Saharan Africa. The reason why a residence permit was issued widely differs by origin. Family migrants more often come from the Maghreb countries (43%), economic migrants from sub-Saharan Africa (53%) and refugees from Asia (40%).

The immigrants' employment rate rose steadily: from 46% in 2010 (67% for men and 26% for women), to 54% in 2011 (74% and 35%) and 59% in 2013 (79% and 42%). In 2010 more than 50% of all immigrants who worked found their current job through relations and recommendation (family, friends, third persons), whereas in 2011 this number decreased to 39% for those people who had a new job in 2011 relative to 2010, replaced by responses to job offers in newspapers, from the *Pôle Emploi* (the French national employment agency), and spontaneous applications. In 2013, 38% of the workers found their current job through relations and recommendations. The numbers of responses to job offers from newspapers and spontaneous job applications increased further. Tables A1 and A2 in the appendix show additional descriptive statistics on socio-demographic characteristics by migration reason, as well as labor market outcomes by gender and year.

25% of the immigrants (equally distributed between men and women) were proposed the language training program, a total of 1479 individuals. Out of a minimum of 0 and a maximum of 400 hours, the average number of prescribed hours is 264, the median number is 200. 73% of all individuals offered a language training program completed their training by 2011. Of those who had completed the training, 95% passed the DILF French language exam (*Diplôme Initial de Langue Française*), which provides them with a recognized diploma. Nonetheless, 70% of the individuals who took the language classes think that this training was not enough to read and speak French – they would have wished to have additional hours.

The compliance rate – i.e., the percentage of immigrants who were assigned language

training and actually attended the classes – was high: 91%. The reason for this high compliance rate is that attending the training sessions of the CAI is a requirement enhancing the probability of renewal of the residence permit. For the individuals who abandoned or did not follow their language training, the most important reasons involved were being pregnant, having to take care of children or health issues. By 2013 almost 90% completed their language training, 97% of whom passed the DILF.

Grouping immigrants by nationality shows that immigrants from Asia are the most likely to be assigned language training (47%), second most likely are immigrants from Europe (33%), whereas immigrants from Maghreb countries, sub-Saharan Africa and other African countries are much less likely to be assigned language training. Australia and Oceania account for the lower range. These numbers may reflect former colonial ties and French language skills before arrival.

Immigrants with lower language skills are less likely to have a job. In 2013, 65% of immigrants with very good French skills had a job, while only 56% of immigrants with little or no French had one. In 2011, 59% of immigrants with very good French had a job, while only 43% of immigrants with little or no French had one. In 2010, the percentages were 48% versus 32%.

#### **4 Regression discontinuity design (RDD)**

As explained, each new legal immigrant to France has to take a language exam assessing her knowledge of the French language, written and spoken. If the result is insufficient, the person is likely to be assigned to a language program. The rule is the following: the treatment (i.e., the linguistic training) is likely to be offered when the test results are inferior or equal to a certain threshold. More precisely, the exam is composed of five parts. The first part is an oral examination, where the individual can obtain a score between zero and 70 points. Only those individuals who have at least 35 points are allowed to take the four additional written tests, with a maximum of 30 possible points. The total score ranges from zero to 100 points and is measured in intervals of five points. 45 is the last failing grade and 50 is the first passing grade. This eligibility rule creates a discontinuity in the relation between the test result and the variables of interest. The discontinuity will be used to estimate the causal effect of the linguistic training, through the method of Regression

## 4.1 Fuzzy RDD

The treatment is heterogeneous in the sense that different individuals can be assigned a different number of hours of language training. The range goes from 0 until 400 hours, in steps of 10. The assignment to the language training and the number of hours do however not solely depend on the fact of passing or not the language exam. The government may take into account other personal and socio-demographic characteristics of the immigrant as well. So, for the same test score, an immigrant coming from a non-francophone country may have a higher probability of being assigned language training (and more hours of classes) than an immigrant from a francophone country (Le Quentrec-Creven, 2014). Nonetheless, the language examination is the most important variable taken into consideration. As Figure 1 shows – considering the total test score of the entry exam as horizontal axis and the number of hours of language training as vertical axis – there is a clear cutoff at the passing threshold.

Given that our running variable is not the only determinant of the treatment status, we employ the method of fuzzy RD design, using the test result to build an instrument for the number of hours of language classes the immigrant was assigned.

Our first and second stage regressions are set up the following way. Equation 1 represents the second stage of the 2SLS specification:

$$Y_i = \beta_0 + \beta_1 LC_i + \beta_2 X_i + \varepsilon_i \quad (1)$$

$Y_i$  is the outcome of interest for individual  $i$  (e.g. labor force participation, employment status, objective language skills, etc.) and  $X_i$  is a vector of control variables. Our main explanatory variable is  $LC_i$ , which is equal to 0 if no language classes were assigned or the number of prescribed hours if the language training was assigned. We are, thus, interested in the coefficient  $\beta_1$ . Equation 2 is the first stage of our specification:

$$LC_i = \alpha_0 + \alpha_1 T_i + \alpha_2 X_i + v_i \quad (2)$$

$T_i$  is the assignment variable, that is a dummy equal to 1 if the immigrant did not succeed

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<sup>8</sup>See Lee and Lemieux (2010) for a survey of RD design in economics and Pinotti (forthcoming) for a recent application of this methodology to assess the effect of immigrants' legal status on crime.

the language test or 0 if she did succeed. More precisely,  $T = 1$  if  $TestResult < cutoff(50)$  and  $T = 0$  if  $TestResult \geq cutoff(50)$ .

The set of control variables of vector  $X$  includes education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the “Ile-de-France” region (i.e. the geographical area around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the immigrant has already spent in France, dummy variables indicating the reason to migrate (the three main options being labor migration, refugee migration, and family migration), a dummy variable equal to 1 if the individual was employed in 2010 (i.e., at the time of the language entry test), and country of origin fixed effects.

## 4.2 Fuzzy RDD with a discrete running variable: two approaches

A peculiarity of our setting concerns the type of running variable. The test result is in intervals of five points, meaning that we consider a discrete assignment variable. Lee and Card (2008) explain that in the case of a discrete running variable, it is not possible to compute and compare average outcomes in a small neighborhood around the threshold. We need, thus, to choose the functional form for the relationship between the treatment variable and the outcomes of interest (Lee and Card, 2008).

For a sufficiently small bandwidth around the threshold we can assume a linear relationship between our outcomes of interest and the treatment variable (Lee and Lemieux, 2010). In this case, we will estimate equations (1) and (2), within a small bandwidth around the cutoff. This approach allows, thus, to use only a subsample of the total number of individuals.

From Lee and Card (2008) we know that a local linear estimator is an imperfect estimator in the case of a discrete running variable. Moreover in our setting we do not have a sufficiently large sample size around the threshold (see Figure 3). Hence, we consider as well a second approach introduced by the same authors, which relies on choosing an appropriate parametric form of the model. This approach has the advantage of using the whole sample, which adds statistical power and robustness to the results.

We choose a vector of polynomials to approximate the functional form of our model. Equations (1) and (2) can be rewritten as follows:

$$Y_i = \beta_0 + \beta_1 LC_i + \beta_2 Dist_i + \beta_3 Dist_i * T_i + \beta_4 X_i + \varepsilon_i \quad (3)$$

$$LC_i = \alpha_0 + \alpha_1 T_i + \alpha_2 Dist_i + \alpha_3 Dist_i * T_i + \alpha_4 X_i + v_i \quad (4)$$

where the variables are defined as in equations (1) and (2).  $Dist_i$  is a row vector of polynomial terms of the test result normalized around its cutoff value, 45 (i.e., the first failing grade). These terms control for the distance from the cutoff and define the functional form of our model.  $Dist_i * T_i$  allows the functional form to vary on each side of the cutoff. For our estimates we will gradually increase the polynomial degree, from the first to the second. We stop at the second order polynomial, following the suggestions of Gelman and Imbens (2016) about the flaws of controlling for higher-order polynomials. We follow the literature in considering  $T$  as the only excluded instrument in equation (4).<sup>9</sup> In all specifications, we cluster standard errors by country of origin and test scores of the initial language exam.

### 4.3 Validity of the RD design

Figure 1 depicts the first stage of our model. The graph plots the relationship between the test score and residuals from a regression of the number of hours of language training on the control variables at baseline. The line shows the predicted values of a local linear smoother with triangular kernel.<sup>10</sup> We can clearly identify a jump around the cutoff: the number of hours of language classes are, indeed, remarkably higher for individuals who failed the test (i.e., who have a test result of 45 or less) with respect to individuals who passed the test.

In Figure 2 we plot labor force participation in 2013 against the running variable (the test result in 2010), again after baseline characteristics have been partialled out. The line shows the predicted values of a local linear smoother with triangular kernel. There is a clear jump around the cutoff: individuals who marginally failed the test, and were thus

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<sup>9</sup>As Angrist and Pischke (2009) suggest, the resulting just-identified IV estimator has the advantage of having good finite sample properties.

<sup>10</sup>In principle we would expect a negative slope of this line both below and above the threshold. A possible reason for the slightly positive slope for test scores larger than the cutoff value is that we condition on a set of control variables at baseline.

more likely to take up the language training program, appear to have a higher probability of participating in the labor force in 2013 than those who marginally passed the test.

Next, in the spirit of the McCrary (2008) test, we plot the density of the forcing variable, to make sure the test scores have not been manipulated around the cutoff. Given that in our context the running variable is discrete, we have checked its density by plotting histograms. Figure 3 shows the density plot over the full sample and around the threshold. Focusing on the density around the threshold, we see that there is a higher mass around the test result of 50, which is the first passing result. However, comparing with the full sample, this is a low density overall: 42 individuals have a test result of 50. Again, these histograms emphasize the importance of the two approaches towards the estimation of the impact of the language classes. The density around the passing threshold is low. Considering the whole sample of individuals – while controlling parametrically for a polynomial of the distance from the threshold – adds statistical power and robustness to the results.

In our context, manipulation by the individuals on the one hand and the provider of the language classes on the other hand, is unlikely.<sup>11</sup> In order for our methodology to be valid, we need to be sure that the individuals close to the cutoff did not purposely fail the language test. This case is, intuitively, to be ruled out, since those who wanted to fail would have failed the test by more than just five or ten points. There is no incentive to try to fail the test marginally – for instance targeting a test score of 40 or 45 –, as this strategy induces the risk of not failing, as well as having a lower number of assigned hours. Also, the test consists of five parts, the individual does not know how many points she gets for each part of the test and the number of points needed to get exactly 45. In addition, the individual does not necessarily know the threshold needed to succeed the entry test. On the other hand, it is equally important that the results were not manipulated by the language class provider. The risk here is that the test scores of individuals who would have marginally failed, were pushed upwards to pass the test, or vice versa. Also this hypothesis can likely be ruled out. As discussed in the previous section, the test score of the entry exam is not the only variable that matters for the assignment to language training. There is no incentive, thus, from the side of the language class provider, to manipulate the test result. If the government wants to assign - or not to assign - language classes to a spe-

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<sup>11</sup>While in this section we explain why manipulation of the assignment variable around the threshold is conceptually unlikely, in Section 5.1 we consider robustness checks to address this potential threat to the validity of our empirical strategy.



cific individual it simply does so. In addition, the test consists of five parts, administered by different people, meaning that one individual examiner can hardly decide by herself whether to fail or pass a person.

Finally, we follow the literature on RD design and carry out a test of balance of pre-treatment characteristics, in order to check whether there is sorting around the threshold. Table 1 shows the reduced form estimates, with 19 different dependent variables in their 2010 values (i.e., observed at the time of the entry test). We regress each one of these variables<sup>12</sup> on the dummy variable equal to 1 if the test result is strictly inferior to 50, including polynomials to the second degree. Ideally, we expect no significant coefficients, which would confirm random sorting around the threshold. And indeed, for all 19 variables we find no significant coefficient, showing no evidence of a discontinuity between test scores and baseline characteristics.

## 5 Results

Table 2 presents our first stage results (Panel A), as well as the second stage estimates (Panel B), using labor force participation as dependent variable.<sup>13</sup> Labor force participation is equal to 1 for all individuals who are employed or registered as unemployed in 2013. By then, almost 90 percent of the individuals have already completed their language training. It is, thus, a point in time that is advanced enough to show first effects on economic integration, but does not allow yet to study long run effects.

Our results are presented in eight columns, where each odd-numbered column shows the estimates without control variables and country of origin fixed effects, and each even-numbered column presents the results from the regressions that include a set of control variables and country of origin fixed effects. The set of control variables is the one presented in Section 4.1.

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<sup>12</sup>Education level, age, time already spent in France, a dummy equal to 1 if the individual is resident in the "Ile-de-France" region, a dummy equal to 1 if the individual is male, a dummy equal to 1 if the person is married, number of children in the household, number of individuals in the household, a dummy equal to 1 if the individual is employed at the time of the entry test, a dummy equal to 1 if the individual is taking other French classes, three dummies for the different migration reasons, and regions of origin. All these variables are measured in 2010.

<sup>13</sup>In this section, the main focus is on labor force participation as dependent variable, because for this outcome of interest we find statistically significant effects of the language classes. In Table 6, we show results from regressions using other labor market outcomes as dependent variables.

The first four columns in the table follow our first approach: local linear regressions around the cutoff. For this approach, in line with the literature on RD design, we determine the optimal bandwidth around the cutoff. For this purpose, we use the optimal bandwidth selector suggested by Calonico, Cattaneo and Titiunik (2014), which is compatible with fuzzy RD designs. The preferred bandwidth includes individuals with test results between 35 and 60 points. For the sake of robustness, we also report the estimation results of a specification using a larger bandwidth, considering individuals with test results between 30 and 65 points.

The second set of results, columns five to eight, follows our second approach, the parametric estimates over the whole sample based on Lee and Card (2008). We increase the polynomial of the forcing variable centered around the threshold (i.e., the distance of the test score from the threshold) from the first to the second degree and include as well interactions with our instrumental variable equal to 1 if the test result is lower than 50.

Looking at the first stage in Table 2, in all specifications we find a highly significant impact of failing the entry test on the number of assigned hours of French language training. The second stage shows the effect of additional 100 hours of training on the probability of participating in the labor force. We observe a significant positive impact, which is robust throughout the specifications. An increase by 100 hours of training raises the probability of participating in the labor force between 14.5 and 26.6 percentage points. This is a sizeable effect, considering the average probability of labor force participation in 2013 being 77%.

Table 3 shows heterogeneity of effects by education level, which is defined as the number of years of education. The range goes from 0 to 17, with a median value of 12, which is equal to a high school degree. The most frequent levels of education are 0 years at school – i.e., no schooling – (26% of the individuals in the sample), 12 years at school (19% of the individuals) and 17 years at school (17% of the individuals). The table shows that the higher the level of education, the bigger the impact of the language training on labor force participation. In fact, for certain specifications we find no statistical significance for individuals with no schooling (see the non-interacted term). Illiterate immigrants have little or no benefit from the language classes, which may suggest the importance of combining language classes and literacy policies for individuals with no schooling. For all our specifications, the coefficients  $a_0$  and  $a_1$  are jointly significant at the median value of education level.

Table 4 analyzes the impact of the language classes on labor force participation by migra-

tion category.<sup>14</sup> Panel A considers labor migration, relative to other types of migration (family, refugee, other). Looking at the coefficients  $a_0$  and  $a_1$  (hours of French classes and the interaction term), as well as the joint p-value, there is a significant higher impact of language classes for labor migrants relative to the other categories. A similar conclusion can be drawn looking at Panel B, where we analyze refugee migration. Although the coefficients are significant and positive, overall the effects for refugees appear to be smaller in magnitude compared to the effects that we find for labor migrants. The coefficients in Panel C, on the other hand, show that family migrants tend to benefit relatively less from the language training in terms of labor force participation, compared to labor migrants and refugees.

In Table 5 we consider labor force participation by gender and age category. Women appear to benefit less from the language training than men. This result is in line with the conclusions of Gathmann (2015), who looks at the impact of waiting time for naturalization on several economic and social integration indicators. She finds that women are more likely to participate in the labor force, but do not invest in education and language skills. In the context of the CAI language training, a large part of the few non-compliers are women. The main reasons for non-compliance among women include being pregnant and having to take care of children at home.

Furthermore, we find that the language training appears more efficient for people below the median age and below the age of 40. In fact, learning a new language might be easier for individuals who arrive in the country at a young age, in line with the findings of Chiswick and Miller (1995).

Table 6 presents regressions using other labor market outcomes as dependent variables. The results for employment, permanent contract, informal work and income per household member are ambiguous. We find results that hint towards a negative impact on the probability of having a full-time job. Although these findings show little robustness throughout the specifications in terms of significance, the negative effect on full-time job may be explained considering the language training as an investment into future employability. Due to the language classes, the individuals might not be able to work full-time, an effect that could last for few years after the classes.

To summarize the main results, we find a strong labor force participation effect without an increase in the probability of employment. However, it is important to stress that our

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<sup>14</sup>In the specifications of Table 4, we compare each type of migration to all other types combined together, because we prefer to estimate a more parsimonious specification with fewer endogenous regressors.

data do not allow to assess the long run effects of the language classes on employment.

### **5.1 Selection into or out of training: Robustness checks for possible manipulation of the assignment variable**

In Section 4.3, we discuss that in principle neither the individuals themselves, nor the language class provider, have an incentive or the power to manipulate the test result around the cutoff. Nonetheless, Figures 2 and 3 show potential issues of selection into or out of training. In Figure 3 we observe a jump in the density of the assignment variable around the cutoff. Notably, there are more individuals who received the first passing grade (50) than individuals who received the first failing grade (45). This may hint towards possible manipulation of the test results by the language class provider (i.e., the government), by increasing the number of immigrants who obtained the first passing grade. The same figure also shows a high density at the "zero" test result and Figure 2 a high value of residuals in the labor force participation equation for observations with a test score equal to "zero". This may indicate a situation of manipulation in which highly motivated individuals hand in a blank exam copy and do not respond to the oral questions, in order to make sure they are assigned language training. To test for the robustness of our results to possible manipulation of the assignment variable, we run two additional sets of regressions.

First, to build an instrumental variable that takes into account possible manipulation by the language class provider, we reclassify certain individuals' test score from 50 to 45. We choose the individuals who marginally passed the test, but are most prone to potentially have been pushed by the government from 45 to 50 points. For this purpose we choose four characteristics that may induce such a positive selection out of training: the education level, the permanence in France (i.e., how long the immigrant has already been in France), the distance of French to the native language (see this variable construction in Section 5.2) and the fact of being employed at the time of the test. Each of these characteristics may be a signal for increased integration potential. For each variable, we calculate the 50th percentile for the individuals from our estimation sample who received 50 points in their test. Then we reclassify the immigrants with values above the 50th percentile from

test score 50 to test score 45.<sup>15</sup>

Denoting with “\*” the “true” unobserved variables and taking education level as an example, we assume that the language class provider manipulated the test score in line with the following misclassification rule:

$$TestResult_i = 50 \quad \text{if} \quad TestResult_i^* = 45 \quad \text{and} \quad EducationLevel_i > 9 \quad (5)$$

$$TestResult_i = TestResult_i^* \quad \text{otherwise} \quad (6)$$

where 9 is the median level of education for individuals with an *observed* test score equal to 50. To build the new instrumental variable, we create a binary variable  $r$  equal to 1 if the individual is reclassified and 0 otherwise. We then generate an IV that takes into account the probability  $p$  of being reclassified:  $T^* - pr$ , where  $T^* = 1$  if  $TestResult^* < cutoff(50)$  and  $T^* = 0$  if  $TestResult^* \geq cutoff(50)$ .<sup>16</sup>

Using this modified instrumental variable in a regression with the *observed* outcome variable as dependent variable, we replicate our main estimates, showing the effect of language classes on labor force participation (i.e., considering specifications similar to those in Table 2). Panel A in Table 7 shows the results following the reclassification according to the education level. The reasoning behind this reclassification comes from a higher language learning potential for higher levels of education. This might induce the language class provider to raise the test result from 45 to 50 for individuals with higher education levels, since they may be considered relatively less in need of language classes. In Panel B, we reclassify individuals according to their permanence in France. The more time immigrants have already spent in France at the time of the entry test, the easier it is for them to learn the new language. For Panel C, we choose to reclassify from 50 to 45 the individuals whose native language has lower distance relative to French. These individuals might need relatively less language training compared to individuals whose native language is more distant from French. Finally, we look at employment at the time of the entry test, in Panel D. The individuals who were employed in 2010 might be considered less vulnerable

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<sup>15</sup>The exception is employment at the time of the test, which is a binary variable. For this variable we reclassify all individuals who were employed in 2010 among those who had a test score equal to 50. In this case, we re-construct our instrument considering that these individuals form 59% of all individuals in our sample with a test result equal to 50.

<sup>16</sup>See Manacorda (2012) for a similar exercise in the context of grade repetition in school using data from Uruguay. In our setting, we exploit the information from government’s documents to choose the criteria that may determine selection out of language classes.

in terms of integration and also less in need of language classes.

Comparing the results with each other and with the results in Table 2, the coefficients remain similar. We can conclude that our estimates are robust to this test. Our findings do not depend on positive selection out of language classes due to manipulation of the test score by the government. In Tables A3 to A7 of the Appendix, we show that findings from this robustness test are similar when we consider estimates on other labor market outcomes.

As a second robustness test, we check whether the high density of "zero" test scores in Figure 3 and the high value of the residuals in the labor force participation equation for immigrants with "zero" test score (see Figure 2) may signal the individuals' manipulation of the assignment variable. Highly motivated individuals may want to participate in free language classes even though their level of French is high enough to achieve the passing grade in the test. We replicate the estimates of specifications similar to those in Table 2, excluding from the estimation sample all the individuals with a "zero" test score. This exercise implicitly considers a worst-case scenario where all immigrants with a "zero" test score cheated their entry test to be assigned to the language classes. Table 8 shows that our results are also robust to this check. In Table A8 of the Appendix, we repeat this robustness test on other labor market outcomes, which appear to be robust as well.

## 5.2 Mechanisms

An important question is to assess the mechanisms through which the language classes impact on economic integration. We point out five possible channels, which we present in Tables 9 and 10.

The first and most direct channel is an objective improvement in the language skills. To check for this channel, we use the results of a written and oral language exam that all individuals in the sample had to take in 2013, as well self-reported language skills. The estimates are presented in Panel A of Table 9. There seems to be no robust significant impact of the language classes on the actual language skills of immigrants. There is a hint towards a negative subjective perception of one's writing skills, which might be related to a negative behavioral effect associated to the disappointment of the absence of impact of the integration policies on employment. This result - whose interpretation seems to be confirmed by other findings below - needs however to be considered with caution. By

2011 over 70% of the language class participants had already completed their training. Measuring language skills in 2013 means quantifying the skills two years after completion of the classes for most individuals. During this time, language skills might have decreased again if no further training was taken. Nonetheless, descriptive statistics show that over 70% of the individuals who were assigned language training do not think the training was enough to learn French, and about the same percentage would have liked to have more hours. Column 3 in Table 9 reports estimates of the heterogeneity of effects, which allow to analyze the role of the distance of the native language to French. To compute measures of language distance, we follow previous literature (for instance, Fearon, 2003; Desmet, Ortuño and Wacziarg, 2012; Alesina, Harnoss and Rapoport, 2016; Laitin and Ramachandran, 2016) and use Ethnologue’s linguistic tree diagrams to determine the number of common nodes between the native language of the immigrants – there are about 100 different native languages in our estimation sample – and French. In particular, we compute language distance  $d$  as follows:

$$d_{French\ Other\ Language} = 1 - \left[ \frac{\# of\ common\ nodes\ between\ French\ and\ Other\ Language}{\frac{1}{2}(\# of\ nodes\ for\ French + \# of\ nodes\ for\ Other\ Language)} \right]^\lambda \quad (7)$$

where in line with previous literature we consider  $\lambda$  equal to 0.5. For instance, French is classified as: Indo-European, Italic, Romance, Italo-Western, Western, Gallo-Iberian, Gallo-Romance, Gallo-Rhaetian, Oil, French. Spanish is instead classified as Indo-European, Italic, Romance, Italo-Western, Western, Gallo-Iberian, Ibero-Romance, West Iberian, Castilian. From equation (7), the two classifications imply that the language distance between French and Spanish is 0.205.

For the objective oral language skills we observe an improvement through language classes for individuals whose native language is not too distant from French (the coefficient is statistically significant up to the 38th percentile of the language distance distribution), while the effect is not statistically significant when computed at the median value of the language distance variable (0.62). The estimates also show a negative effect of the language classes on subjective writing skills: this coefficient is significant at the median value of language distance.

A second possible channel works through a signaling effect. Potential employers could see a language diploma as a signal for language proficiency and higher ability of the immigrant. Indeed, after completion of the number of assigned hours, the participants take

the DILF exam, which provides them with a diploma recognizing the acquisition of basic French skills. By 2013, 97% of the participants passed this test, which hints at actual efficiency of the language training. If the signaling effect were an important driver of our main findings, we would expect to observe a significant increase in the employment rate. However, as we have already discussed, Table 6 shows that this is not the case. This finding presents evidence against a signaling effect at work.

The third channel we look at is based on network effects. There is literature showing the role of networks (especially high quality networks) in the integration process (e.g. McKenzie and Rapoport, 2010). The language training could give access to new networks of French friends, international friends or friends from the same country of origin. For instance, this may happen if the individual improves her language skills and becomes more sociable. Also, during the period of the language classes, the immigrant might interact more with her classmates, while having less time to socialize with other people. Panel B in Table 9 shows the estimates for the probability of having new friends in 2011 and 2013. We find some evidence of a negative impact on making new French friends in 2011 and no impact on other potential networks.

As a fourth channel, we consider behavioral effects. In principle, we would expect positive effects if immigrants are encouraged to participate in the labor force because the government implements an integration plan. Instead, we find suggestive evidence of negative behavioral effects, probably because immigrants are demotivated after they realize the integration plan is not effective in increasing the probability of employment. Indicators for this channel are the lack of an impact on self-reported language skills (and the negative effect on subjective perception of writing skills), as well as a negative effect of the language classes on "feeling at home" in France, interest in French politics and interest in the origin country's politics (see Panel B of Table 10).

Finally, we find evidence of a positive effect of the training through our fifth channel, the channel of information. The participants to the training are likely to benefit from positive interactions that the language classes produce. Before, during and after the classes the participants can use their time to exchange about their experiences in France and to give each other important advice on the French life and labor market. This exchange involves not only the participants, but also the teacher. For immigrants, it is simple and time-saving to talk to the teacher and other participants in class, as compared to trying to find the information on their own elsewhere. This is how attendance to the classes can



provide a valuable source of information on life in France and job search strategies.<sup>17</sup> In Panel A of Table 10, outcome variables related to this channel – such as searching and / or having found a job through the *Pôle Emploi* (the French national employment agency), having passed the drivers' license in France and having asked for the recognition of the latest academic diploma – all show positive significant coefficients.

## 6 Concluding remarks

We examine the impact of language training on the economic integration of immigrants in France. The language classes are offered by the French Ministry of the Interior after the immigrant signs the *Contrat d'accueil et d'intégration*. The training is more likely to be available when the test score of an initial language exam is below a certain threshold.

Our study is novel and important in at least two directions. First, we use a local randomized experiment in the form of a Regression Discontinuity Design to estimate the effects of language training on immigrants' integration. Second, we look at the mechanisms through which language classes impact on a set of outcome variables related to economic integration.

We find that the number of hours of training significantly increases labor force participation. The effect appears to be smaller for women, individuals above the median age, and family migrants relative to labor migrants and refugees. The language training shows a significantly higher impact on labor force participation for individuals with higher levels of education. Furthermore, the language classes do not appear to increase the probability of employment. The latter finding however needs to be interpreted with caution because our data do not allow to analyze long run effects.

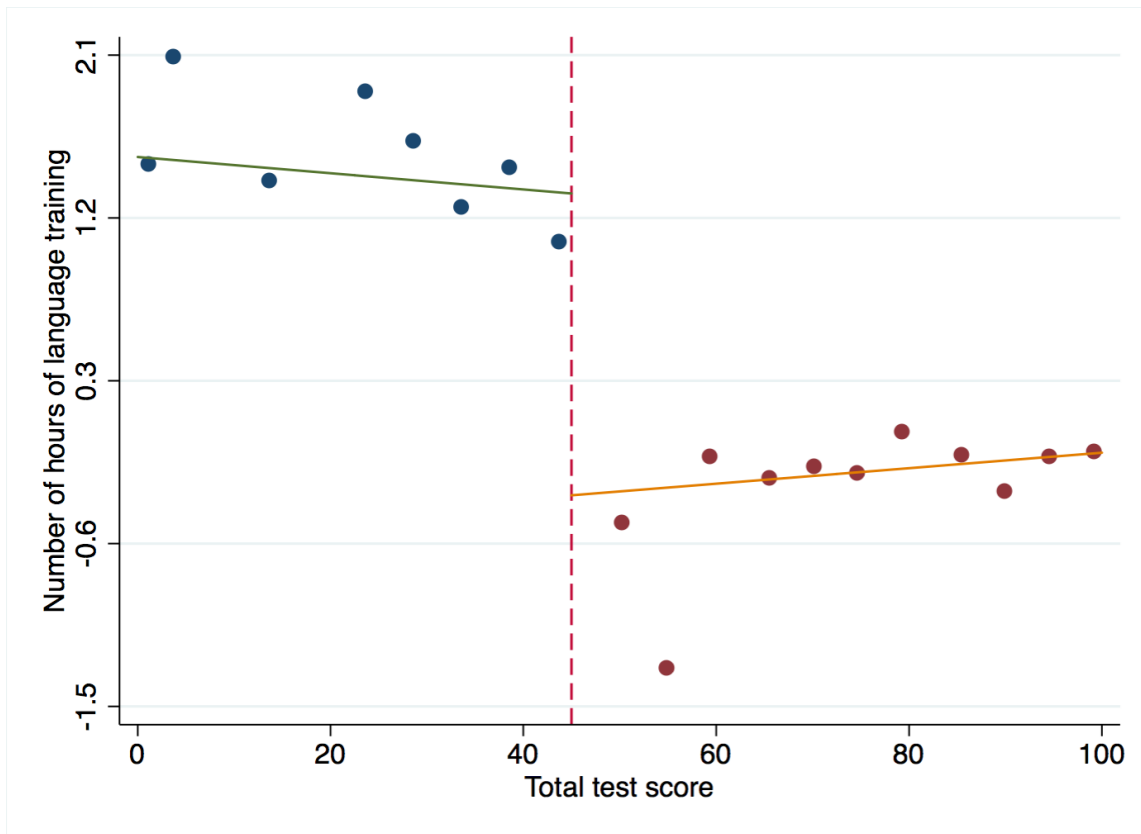
Finally, we explore different mechanisms that explain the effect of language classes on labor force participation. We find little evidence for an objective improvement of language skills, for a signaling effect of the diploma received after the end of the language classes, and for an increase in the size of the network. Instead, we find negative behavioral effects that may derive from disappointment from the integration plan. Finally, the information on job search strategies that individuals derive from the time spent with their classmates

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<sup>17</sup>The *Contrat d'accueil et d'intégration* also includes an information session on life in France, which lasts 6 hours only. The advantage of the interactions during the language classes is that the average time length is higher, and therefore the immigrant has more time to think and ask about information that is useful for her integration.

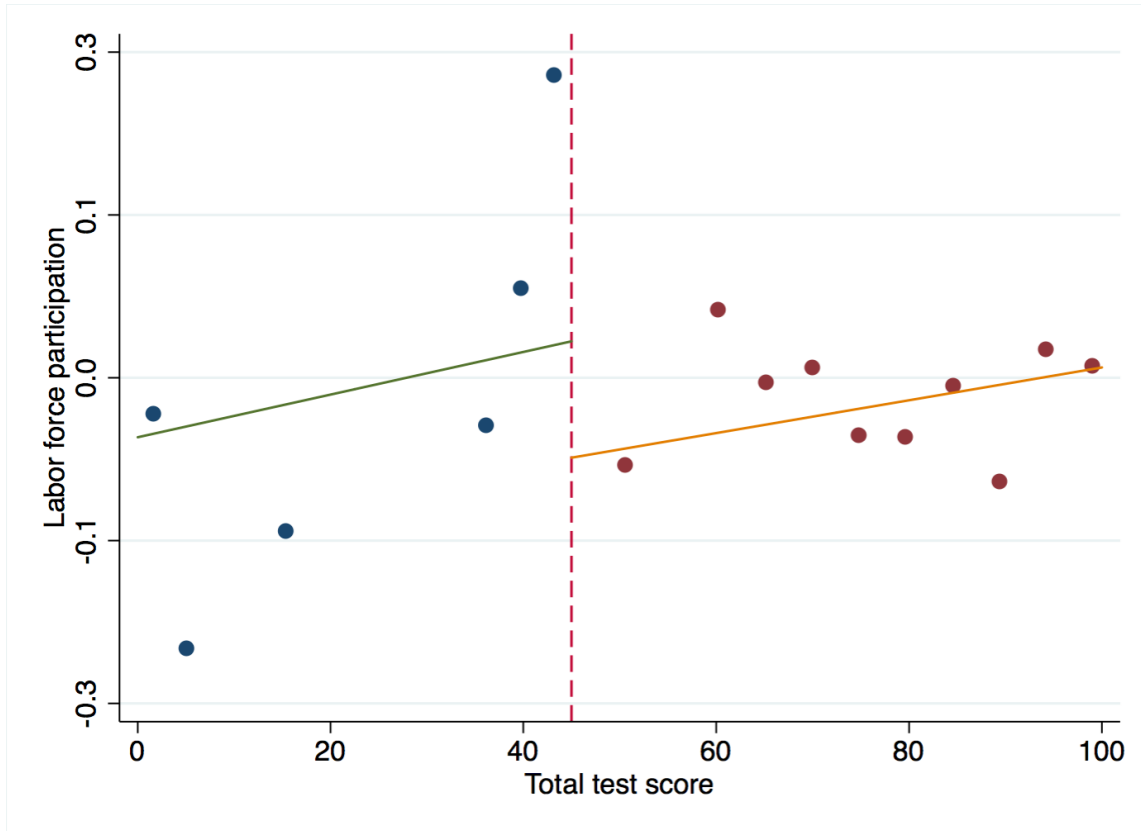
and teachers during the classes appear to be an important channel that helps immigrants to integrate in the French labor market.

Figure 1: First stage graph



First stage plot. This figure plots residuals from a regression of "Number of hours of language training" on a set of covariates (education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the "Ile-de-France" region (i.e. the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the immigrant has already spent in France, dummy variables indicating the reason to migrate, a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects). The variable "Total test score" refers to the language entry exam and is observed in interval of 5 points. The line shows the predicted values of a local linear smoother with triangular kernel.

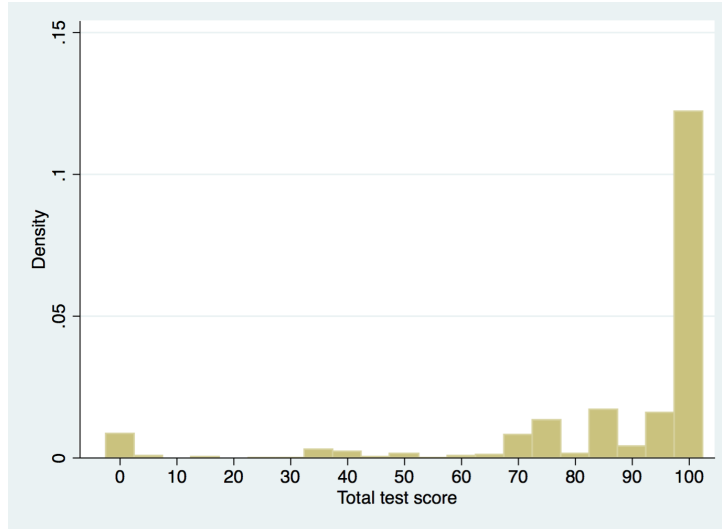
Figure 2: Reduced form graph



Reduced-form relationship. This figure plots residuals from a regression of "Labor force participation" (observed in 2013) on a set of covariates from the baseline survey (education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the "Ile-de-France" region (i.e., the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the immigrant has already spent in France, dummy variables indicating the reason to migrate, a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects). The variable "Total test score" refers to the language entry exam and is observed at baseline in interval of 5 points. The line shows the predicted values of a local linear smoother with triangular kernel.

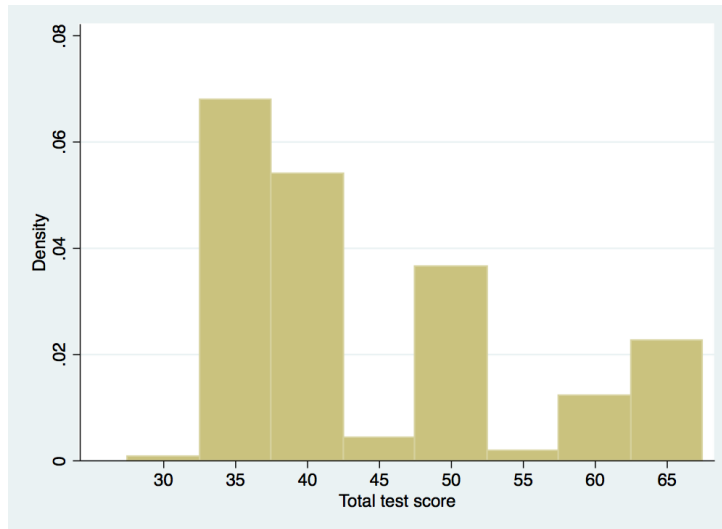
Figure 3:

Density plot over the whole sample



The y-axis represents the density, the x-axis the total test score (which refers to the language entry exam and is observed at baseline in interval of 5 points).

Density plot zoomed on the threshold



The y-axis represents the density, the x-axis the total test score (which refers to the language entry exam and is observed at baseline in interval of 5 points).

Table 1: Test of Balance of Pre-Treatment Characteristics, in 2010.

**Test of Balance of Pre-Treatment Characteristics, in 2010.**

<b>Panel A: Demographic and household characteristics</b>										
	Education level	Age	Permanence in France	Ile-de-France	Married	Male	Number of children	Number of people in HH	Employed in 2010	Taking other French classes
I(test score initial language exam < 50)	-0.772 [2.384]	0.787 [4.099]	-2.071 [1.866]	-0.397 [0.243]	-0.048 [0.125]	0.152 [0.211]	0.013 [0.379]	-0.897 [0.593]	-0.205 [0.202]	0.010 [0.136]
Polynomial of degree	2	2	2	2	2	2	2	2	2	2
Observations	4,969	5,299	5,298	5,307	4,231	5,307	5,307	5,307	5,303	5,304
<b>Panel B: Types of migration and regions of origin</b>										
	Work migration	Refugee migration	Family migration	Europe	Asia	Maghreb	Sub-Sah. Africa	Africa (other)	America & Oceania	
I(test score initial language exam < 50)	-0.080 [0.097]	0.120 [0.185]	0.035 [0.217]	0.113 [0.171]	-0.256 [0.271]	0.315 [0.452]	0.019 [0.219]	-0.098 [0.159]	-0.093 [0.118]	
Polynomial of degree	2	2	2	2	2	2	2	2	2	
Observations	5,307	5,307	5,307	5,307	5,307	5,307	5,307	5,307	5,307	

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.

The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 2, interacted with I(test score initial language exam < 50).

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 2: Assigned French Language Training and Labor Force Participation.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Local linear estimates		Local linear estimates		Parametric estimates			
	[35, 60]	[35, 60]	[30, 65]	[30, 65]	[30, 65]			
<b>Panel A: First stage regressions (dependent variable: number of assigned hours of French language training/100)</b>								
(test score initial language exam < 50)	3.317*** [0.902]	3.294*** [0.936]	3.154*** [0.889]	2.910*** [0.652]	1.591*** [0.312]	1.790*** [0.287]	2.665*** [0.499]	2.576*** [0.466]
<b>Panel B: Second stage regressions (dependent variable: labor force participation dummy)</b>								
Hours of French classes (/100)	0.184** [0.084]	0.266* [0.141]	0.167** [0.081]	0.190* [0.106]	0.047 [0.062]	0.145** [0.072]	0.174** [0.077]	0.206** [0.098]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	67	67	76	76	2,168	2,168	2,168	2,168

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam. The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with (test score initial language exam < 50). The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 3: Assigned French Language Training and LFP. Heterogeneity effects by education level.

	(1)	(2)	(3)	(4)
	Local linear estimates		Parametric estimates	
	[35, 60]	[30, 65]	[30, 65]	
<b>Assigned French Language Training and Labor Force Participation. Heterogeneity effects by education level.</b>				
<b>Panel A: Second stage regressions (dependent variable: labor force participation dummy)</b>				
Hours of French classes (/100) (a <sub>0</sub> )	0.095 [0.072]	0.097 [0.065]	0.110* [0.066]	0.162* [0.088]
Hours of French classes (/100) X education level (a <sub>1</sub> )	0.024* [0.013]	0.024 [0.015]	0.008* [0.004]	0.011* [0.006]
H <sub>0</sub> : a <sub>0</sub> +12*a <sub>1</sub> =0 (p-value)	0.0238	0.0631	0.0419	0.0363
Polynomial of degree	1	1	1	2
Control variables	Yes	Yes	Yes	Yes
Country of origin fixed effects	Yes	Yes	Yes	Yes
Observations	67	76	2,168	2,168

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam. The median value of education level is 12 years. The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with (test score initial language exam < 50). The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 4: Assigned French Language Training and LFP. Heterogeneity of effects by migration category.

	(1)	(2)	(3)	(4)
	Local linear estimates		Parametric estimates	
<b>Assigned French Language Training and Labor Force Participation. Heterogeneity of effects by migration category (labor migration, refugee migration, and family migration).</b>				
<b>Panel A: Second stage regressions (dependent variable: labor force participation dummy). Labor migration versus other types of migration (refugee and family migration)</b>				
Hours of French classes (/100) ( $a_0$ )	[35, 60] 0.268*	[30, 65] 0.190*	0.144**	0.205**
	[0.147]	[0.106]	[0.072]	[0.098]
Hours of French classes (/100) X labor migration ( $a_1$ )	0.360	0.013	0.076	0.095
	[1.363]	[0.701]	[0.052]	[0.070]
labor migration	-0.003	0.016	-0.031**	-0.034**
	[0.514]	[0.233]	[0.015]	[0.016]
$H_0: a_0 + a_1 = 0$ (p-value)	0.6581	0.7724	0.0251	0.0240
<b>Panel B: Second stage regressions (dependent variable: labor force participation dummy). Refugee migration versus other types of migration (labor and family migration)</b>				
Hours of French classes (/100) ( $a_0$ )	0.266*	0.198	0.142*	0.204**
	[0.149]	[0.126]	[0.075]	[0.103]
Hours of French classes (/100) X refugee migration ( $a_1$ )	0.008	-0.039	0.026	0.010
	[0.162]	[0.145]	[0.058]	[0.064]
refugee migration	-0.009	-0.012	0.002	0.006
	[0.393]	[0.367]	[0.047]	[0.049]
$H_0: a_0 + a_1 = 0$ (p-value)	0.0832	0.1175	0.0158	0.0164
<b>Panel C: Second stage regressions (dependent variable: labor force participation dummy). Family migration versus other types of migration (labor and refugee migration)</b>				
Hours of French classes (/100) ( $a_0$ )	0.409*	0.199	0.176**	0.232**
	[0.245]	[0.127]	[0.068]	[0.092]
Hours of French classes (/100) X family migration ( $a_1$ )	-0.153	-0.011	-0.037	-0.031
	[0.204]	[0.132]	[0.045]	[0.049]
family migration	-0.210	-0.199	0.021	0.015
	[0.337]	[0.282]	[0.051]	[0.052]
$H_0: a_0 + a_1 = 0$ (p-value)	0.0740	0.1050	0.0648	0.0478
Polynomial of degree	1	1	1	2
Control variables	Yes	Yes	Yes	Yes
Country of origin fixed effects	Yes	Yes	Yes	Yes
Observations	67	76	2,168	2,168

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam. The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with (test score initial language exam < 50). The set of control variables includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1



Table 5: Assigned French Language Training and LFP. Heterogeneity of effects by gender and age.

	(1)	(2)	(3)	(4)
	Local linear estimates		Parametric estimates	
	[35, 60]	[30, 65]		
<b>Assigned French Language Training and Labor Force Participation. Heterogeneity of effects by gender and age.</b>				
<b>Panel A: Second stage regressions (dependent variable: labor force participation dummy). Women and Men</b>				
Hours of French classes (/100) (a <sub>0</sub> )	0.575*	0.266*	0.153**	0.213**
	[0.322]	[0.145]	[0.072]	[0.104]
Hours of French classes (/100) X female (a <sub>1</sub> )	-0.401*	-0.142	-0.014	-0.012
	[0.237]	[0.123]	[0.041]	[0.045]
female	-0.081	-0.203	-0.187***	-0.189***
	[0.262]	[0.154]	[0.030]	[0.030]
H <sub>0</sub> : a <sub>0</sub> +a <sub>1</sub> =0 (p-value)	0.2833	0.2619	0.0846	0.0466
<b>Panel B: Second stage regressions (dependent variable: labor force participation dummy). Above and below median age (32 years old)</b>				
Hours of French classes (/100) (a <sub>0</sub> )	0.349*	0.250*	0.168**	0.233**
	[0.184]	[0.134]	[0.075]	[0.098]
Hours of French classes (/100) X above median age (a <sub>1</sub> )	-0.110	-0.097	-0.055	-0.059*
	[0.115]	[0.087]	[0.037]	[0.035]
H <sub>0</sub> : a <sub>0</sub> +a <sub>1</sub> =0 (p-value)	0.1068	0.1631	0.0712	0.0545
<b>Panel C: Second stage regressions (dependent variable: labor force participation dummy). Top quintile age (&gt;=40 years old)</b>				
Hours of French classes (/100) (a <sub>0</sub> )	0.266*	0.195*	0.156**	0.214**
	[0.140]	[0.108]	[0.074]	[0.099]
Hours of French classes (/100) X top quintile age (a <sub>1</sub> )	0.001	-0.030	-0.053	-0.053
	[0.101]	[0.092]	[0.043]	[0.046]
H <sub>0</sub> : a <sub>0</sub> +a <sub>1</sub> =0 (p-value)	0.0990	0.1955	0.1264	0.0793
Polynomial of degree	1	1	1	2
Control variables	Yes	Yes	Yes	Yes
Country or origin fixed effects	Yes	Yes	Yes	Yes
Observations	67	76	2,168	2,168

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with [(test score initial language exam < 50).  
 The set of control variables includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region 'Ile-de-France' (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 6: Assigned French Language Training and Other Measures of Labor Market Integration.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Assigned French Language Training and other measures of Labor Market Integration.							
		Local linear estimates				Parametric estimates		
	[35, 60]	[35, 60]	[30, 65]	[30, 65]				
<b>Panel A: Dependent variable: employment (conditional on labor force participation). Second stage regressions.</b>								
Hours of French classes (/100)	-0.041 [0.122]	-0.134** [0.056]	-0.043 [0.121]	-0.069 [0.066]	-0.015 [0.068]	0.053 [0.041]	-0.087 [0.105]	0.007 [0.068]
Observations	51	51	58	58	1,734	1,734	1,734	1,734
<b>Panel B: Dependent variable: full-time work (conditional on employment). Second stage regressions.</b>								
Hours of French classes (/100)	-0.126 [0.204]	0.557 [0.486]	-0.184 [0.206]	0.976 [0.874]	-0.317** [0.143]	-0.197*** [0.067]	-0.263 [0.200]	-0.240** [0.120]
Observations	46	46	52	52	1,387	1,387	1,387	1,387
<b>Panel C: Dependent variable: permanent contract (conditional on employment). Second stage regressions.</b>								
Hours of French classes (/100)	-0.173 [0.210]	0.256 [0.282]	-0.120 [0.191]	0.271 [0.376]	-0.232* [0.118]	-0.059 [0.065]	-0.116 [0.132]	0.084 [0.117]
Observations	46	46	52	52	1,387	1,387	1,387	1,387
<b>Panel D: Dependent variable: informal work (conditional on employment). Second stage regressions.</b>								
Hours of French classes (/100)	-0.090 [0.092]	0.055 [0.189]	-0.143* [0.079]	-0.010 [0.206]	0.029 [0.065]	0.004 [0.031]	-0.090 [0.074]	-0.053 [0.055]
Observations	46	46	52	52	1,387	1,387	1,387	1,387
<b>Panel E: Dependent variable: income per household individual. Second stage regressions.</b>								
Hours of French classes (/100)	-56.511 [106.643]	184.562** [89.804]	-58.650 [87.648]	116.946 [82.467]	-62.503 [86.010]	-100.908** [51.142]	-148.735* [87.866]	-50.056 [54.728]
Observations	61	61	69	69	2,040	2,040	2,040	2,040
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with (test score initial language exam < 50).  
 The set of control variables includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Île-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*- p<0.01, \*- p<0.05, \* p<0.1

Table 7: Positive Selection out of Language Training: Labor Force Participation.

Positive Selection Out Of Language Training: Labor Force Participation.								
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	[35, 60]	Local linear estimates [35, 60]		[30, 65]		Parametric estimates		
<b>Panel A: By education level.</b>								
<b>Second stage regressions (dependent variable: labor force participation dummy)</b>								
Hours of French classes (/100)	0.168** [0.073]	0.215** [0.098]	0.153** [0.071]	0.144* [0.077]	0.040 [0.063]	0.143** [0.071]	0.160** [0.071]	0.188** [0.095]
<b>Panel B: By permanence in France.</b>								
<b>Second stage regressions (dependent variable: labor force participation dummy)</b>								
Hours of French classes (/100)	0.198** [0.090]	0.245* [0.135]	0.209** [0.099]	0.171 [0.112]	0.063 [0.063]	0.153** [0.072]	0.219** [0.086]	0.213** [0.105]
<b>Panel C: By language distance (Francophone vs. Non-Francophone).</b>								
<b>Second stage regressions (dependent variable: labor force participation dummy)</b>								
Hours of French classes (/100)	0.174** [0.082]	0.264* [0.140]	0.151* [0.080]	0.200* [0.111]	0.039 [0.064]	0.143** [0.073]	0.157** [0.079]	0.199** [0.101]
<b>Panel D: By employment in 2010.</b>								
<b>Second stage regressions (dependent variable: labor force participation dummy)</b>								
Hours of French classes (/100)	0.183** [0.083]	0.235* [0.121]	0.195** [0.092]	0.187* [0.106]	0.061 [0.062]	0.150** [0.072]	0.208** [0.082]	0.208** [0.102]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	67	67	76	76	2,168	2,168	2,168	2,168

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 8: Assigned French Language Training and LFP: Excluding Test Result = 0.

Assigned French Language Training and Labor Force Participation: Excluding Test Result = 0.				
	(1)	(2)	(3)	(4)
		Parametric estimates		
<b>Second stage regressions (dependent variable: labor force participation dummy)</b>				
Hours of French classes (/100)	0.085 [0.059]	0.178** [0.070]	0.160** [0.072]	0.181* [0.097]
Polynomial of degree	1	1	2	2
Control variables	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes
Observations	2,093	2,093	2,093	2,093

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 9: Mechanisms. Improvement in Language Skills and Networks.

Mechanisms. Improvement in language skills and networks.	(1)	(2)	(3)
	Parametric estimates		
<b>Panel A: Improvement in language skills</b>			
<b>Dependent variable: oral skills. Second stage regressions.</b>			
Hours of French classes (/100) (a <sub>0</sub> )	0.403 [0.749]	0.526 [0.589]	2.260** [1.094]
Hours of French classes (/100) X language distance (a <sub>1</sub> )			-2.095*** [0.745]
H <sub>0</sub> : a <sub>0</sub> +0.62*a <sub>1</sub> =0 (p-value)			0.1925
Observations	2,018	2,018	2,000
<b>Dependent variable: writing skills. Second stage regressions.</b>			
Hours of French classes (/100) (a <sub>0</sub> )	-1.138 [2.239]	-1.005 [2.049]	-1.006 [3.130]
Hours of French classes (/100) X language distance (a <sub>1</sub> )			-0.301 [2.035]
H <sub>0</sub> : a <sub>0</sub> +0.62*a <sub>1</sub> =0 (p-value)			0.5911
Observations	1,915	1,915	1,903
<b>Dependent variable: subjective oral skills. Second stage regressions.</b>			
Hours of French classes (/100) (a <sub>0</sub> )	0.144 [0.099]	0.087 [0.104]	0.167 [0.143]
Hours of French classes (/100) X language distance (a <sub>1</sub> )			-0.101 [0.089]
H <sub>0</sub> : a <sub>0</sub> +0.62*a <sub>1</sub> =0 (p-value)			0.3444
Observations	2,163	2,163	2,136
<b>Dependent variable: subjective writing skills. Second stage regressions.</b>			
Hours of French classes (/100) (a <sub>0</sub> )	-0.112 [0.100]	-0.222* [0.114]	-0.272* [0.140]
Hours of French classes (/100) X language distance (a <sub>1</sub> )			0.062 [0.123]
H <sub>0</sub> : a <sub>0</sub> +0.62*a <sub>1</sub> =0 (p-value)			0.0395
Observations	2,170	2,170	2,142
<b>Dependent variable: individual takes other French classes. Second stage regressions.</b>			
Hours of French classes (/100)	0.017 [0.035]	0.014 [0.030]	
Observations	2,170	2,170	
<b>Dependent variable: interview held in French. Second stage regressions.</b>			
Hours of French classes (/100)	0.082 [0.099]	-0.004 [0.080]	
Observations	2,170	2,170	
Polynomial of degree	2	2	2
Control variables	No	Yes	Yes
Country of origin fixed effects	No	Yes	Yes
<b>Panel B: Networks</b>			
<b>Dependent variable: dummy equal to 1 if new French friends in 2013. Second stage regressions.</b>			
Hours of French classes (/100)	0.013 [0.027]	-0.008 [0.023]	
Observations	1,642	1,642	
<b>Dependent variable: dummy equal to 1 if new international friends in 2013. Second stage regressions.</b>			
Hours of French classes (/100)	-0.036 [0.092]	-0.050 [0.076]	
Observations	1,642	1,642	
<b>Dependent variable: dummy equal to 1 if new friends from the same country of origin in 2013. Second stage regressions.</b>			
Hours of French classes (/100)	0.023 [0.088]	0.058 [0.076]	
Observations	1,642	1,642	
<b>Dependent variable: dummy equal to 1 if new French friends in 2011. Second stage regressions.</b>			
Hours of French classes (/100)	-0.072* [0.040]	-0.069* [0.039]	
Observations	1,665	1,665	
<b>Dependent variable: dummy equal to 1 if new international friends in 2011. Second stage regressions.</b>			
Hours of French classes (/100)	0.061 [0.096]	-0.024 [0.070]	
Observations	1,665	1,665	
<b>Dependent variable: dummy equal to 1 if new friends from the same country of origin in 2011. Second stage regressions.</b>			
Hours of French classes (/100)	0.011 [0.087]	0.093 [0.062]	
Observations	1,665	1,665	
Polynomial of degree	2	2	
Control variables	No	Yes	
Country of origin fixed effects	No	Yes	

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam. The median value of language distance is 0.62. The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I (test score initial language exam < 50). The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other; Family migration is the reference category), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 10: Mechanisms. Information Effect and Behavioral Effect.

	(1)	(2)
	Parametric estimates	
<b>Mechanisms. Information effect and behavioral effect.</b>		
<b>Panel A: Information effect</b>		
<b>Dependent variable: Job search trough "Pole emploi". Second stage regressions.</b>		
Hours of French classes (/100)	0.159** [0.039]	0.140** [0.027]
Observations	363	363
<b>Dependent variable: Found his job through "Pole emploi". Second stage regressions.</b>		
Hours of French classes (/100)	0.340 [0.214]	0.255* [0.146]
Observations	1,344	1,344
<b>Dependent variable: Passed driver's license in France. Second stage regressions.</b>		
Hours of French classes (/100)	0.176** [0.069]	0.195*** [0.074]
Observations	797	797
<b>Dependent variable: Applied for recognition of the latest academic diploma. Second stage regressions.</b>		
Hours of French classes (/100)	0.353* [0.198]	0.640** [0.273]
Observations	665	665
<b>Panel B: Behavioral effect</b>		
<b>Dependent variable: Feels "at home" in France. Second stage regressions.</b>		
Hours of French classes (/100)	-0.326* [0.178]	-0.311** [0.157]
Observations	2,157	2,157
<b>Dependent variable: Interested in French politics. Second stage regressions.</b>		
Hours of French classes (/100)	-0.532** [0.216]	-0.395** [0.186]
Observations	2,147	2,147
<b>Dependent variable: Interested in European politics. Second stage regressions.</b>		
Hours of French classes (/100)	-0.337 [0.224]	-0.217 [0.169]
Observations	2,141	2,141
<b>Dependent variable: Interested in origin country politics. Second stage regressions.</b>		
Hours of French classes (/100)	-0.591** [0.238]	-0.379* [0.216]
Observations	2,148	2,148
Polynomial of degree	2	2
Control variables	No	Yes
Country of origin fixed effects	No	Yes

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam. The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with (test score initial language exam < 50). The set of control variables includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

## References

- [1] Algan, Y., Dustman, C., Glitz, A. and A. Manning (2010) "The Economic Situation of First and Second Generation Immigrants in France, Germany and UK", *Economic Journal*, Vol. 120(542), pp. F4–F30.
- [2] Alesina, A., Harnoss, J. and H. Rapoport (2016) "Birthplace diversity and economic prosperity", *Journal of Economic Growth*, Vol. 21(2), pp. 101-138.
- [3] Angrist, J., Battistin, E. and D. Vuri (2016) "In a Small Moment: Class Size and Moral Hazard in the Mezzogiorno", *American Economic Journal: Applied Economics*, forthcoming.
- [4] Angrist, J. and J.-S. Pischke (2009) "Mostly Harmless Econometrics: An Empiricist's Companion", Princeton University Press, 2009.
- [5] Borjas, G. J. (1994) "The Economics of Immigration", *Journal of Economic Literature*, Vol. XXXII, pp. 1667-1717.
- [6] Bleakley, H. and A. Chin (2004) "Language Skills and Earnings: Evidence from Childhood Immigrants", *Review of Economics and Statistics*, Vol. 86, pp. 481-496.
- [7] Bleakley, H. and A. Chin (2010) "Age at Arrival, English Proficiency, and Social Assimilation Among US Immigrants", *American Economic Journal: Applied Economics*, Vol. 2(1), pp. 165–192.
- [8] Borjas, G. J. (1999) "The economic analysis of immigration.", In *Handbook of labor economics*, Vol. 3, ed. Orley Ashenfelter and David Card, pp. 1697–1760. Amsterdam: Elsevier.
- [9] Calonico, S., Cattaneo M. D. and R. Titiunik (2014) "Robust data-driven inference in the regression-discontinuity design", *The Stata Journal*, Vol. 14(4), pp. 909-946.
- [10] Cappellari, L. and A. Di Paolo (2015) "Bilingual Schooling and Earnings: Evidence from a Language-in-Education Reform", *IZA Discussion Papers*, N. 9431, Institute for the Study of Labor (IZA).
- [11] Cattaneo, M. D., Keele, L., Titiunik, R. and G. Vazquez-Bare (2015) "Identification in Regression Discontinuity Designs with Multiple Cutoffs", *Work in progress*, University of Michigan
- [12] Chin, A., Daysal, N. M., and S. A. Imberman (2013) "Impact of bilingual education programs on limited English proficient students and their peers: Regression discontinuity evidence from Texas", *Journal of Public Economics*, Vol. 107, pp. 63-78.
- [13] Chiswick, B. R. (1991) "Speaking, reading, and earnings among low-skilled immigrants" *Journal of Labor Economics*, Vol. 9(2), pp. 149-170.
- [14] Chiswick, B. R., and P. W. Miller (1995) "The endogeneity between language and earnings: International analyses." *Journal of Labor Economics*, Vol. 13(2), pp. 246-288.
- [15] Desmet, K., Ortuño, I. and R. Wacziarg (2012) "The Political Economy of Linguistic Cleavages", *Journal of Development Economics*, Vol. 97, pp. 322-338.

- [16] Direction générale des étrangers en France, Département des statistiques, des études et de la documentation (2010) "Enquete Longitudinale sur l'Integration des Primo-Arrivants 2010 – Vague 1: Documentation et Guide d'utilisation"
- [17] Direction générale des étrangers en France, Département des statistiques, des études et de la documentation (2011) "Enquete Longitudinale sur l'Integration des Primo-Arrivants 2011 – Vague 2: Documentation et Guide d'utilisation"
- [18] Direction générale des étrangers en France, Département des statistiques, des études et de la documentation (2013) "Enquete Longitudinale sur l'Integration des Primo-Arrivants 2013 – Vague 3: Documentation et Guide d'utilisation"
- [19] Dustmann, C. and F. Fabbri (2003) "Language proficiency and labour market performance of immigrants in the UK", *Economic Journal*, Vol. 113(489), pp. 695–717.
- [20] Dustmann, C., Fasani, F. and B. Spedale (forthcoming) "Illegal migration and consumption behavior of immigrant households", *Journal of the European Economic Association*.
- [21] Dustmann, C. and A. van Soest (2001) "Language fluency and earnings: Estimation with misclassified language indicators", *Review of Economics and Statistics*, Vol. 83(4), pp. 663–674.
- [22] Fearon, J. D. (2003) "Ethnic and cultural diversity by country", *Journal of Economic Growth*, Vol. 8(2), pp. 195–222.
- [23] Gathmann, C. (2015) "Naturalization and citizenship: Who benefits?", *IZA World of Labor, Institute for the Study of Labor (IZA)* February, p. 125.
- [24] Gelman, A. and G. Imbens (2016) "Why High-order Polynomials Should not be Used in Regression Discontinuity Designs", Mimeo
- [25] Hinnerich, B. T. and P. Pettersson-Libdom (2014) "Democracy, Redistribution, and Political Participation: Evidence from Sweden 1919-1938", *Econometrica*, Vol. 82(3), pp. 961-993.
- [26] Jacob, B. A. and L. Lefgren (2004) "Remedial Education and Student Achievement: A Regression-Discontinuity Analysis", *The Review of Economics and Statistics*, Vol. 86(1), pp. 226-244.
- [27] Kossoudji, S. A. and D. Cobb-Clark (2002) "Coming Out of the Shadows: Learning about Legal Status and Wages from the Legalized Population", *Journal of Labor Economics*, Vol. 20, pp. 598-628.
- [28] Laitin, D. D. and R. Ramachandran (2016) "Language Policy and Human Development", *American Political Science Review*, 110(3), pp. 457-480.
- [29] Landsberger, H. A. (1958) *Hawthorne Revisited. Management and the worker: its critics, and developments in human relations in industry.*, Ithaca, New York: Cornell University, 132 pp.
- [30] Le Quentrec-Creven, G. (2014) "De la formation linguistique à l'intégration des nouveaux-migrants ?", Actes du XVIIe colloque international de l'AIDELF sur *Démographie et politiques sociales*, Ouagadougou; pp. 1-22.

- [31] Lee, D. S., and D. Card (2008) "Regression discontinuity inference with specification error", *Journal of Econometrics*, Vol. 142, pp. 655-674.
- [32] Lee, D. and T. Lemieux (2010) "Regression Discontinuity Designs in Economics", *Journal of Economic Literature*, Vol. 48(2), pp. 281-355.
- [33] Manacorda, M. (2012) "The Cost of Grade Retention", *The Review of Economics and Statistics*, Vol. 94(2), pp. 596-606.
- [34] Matsudaira, J. D. (2008) "Mandatory summer school and student achievement", *Journal of Econometrics*, Vol. 142, pp. 829-850.
- [35] McCrary, J. (2008) "Manipulation of the running variable in the regression discontinuity design: A density test.", *Journal of Econometrics*, Vol. 142(2), pp. 698-714.
- [36] McKenzie, D. and H. Rapoport (2010) "Self-Selection Patterns in Mexico-U.S. Migration: The Role of Migration Networks.", *The Review of Economics and Statistics*, Vol. 92(4), pp. 811-821.
- [37] Pinotti, P. (forthcoming) "Clicking on Heaven's Door: The Effect of Immigrant Legalization on Crime", *American Economic Review*.
- [38] Sarvimäki, M. and K. Hämäläinen (2016) "Integrating Immigrants: The Impact of Restructuring ALMP", *Journal of Labor Economics*, Vol. 34(2), pp. 479-508.
- [39] Urquiola, M. and E. Verhoogen (2009) "Class-Size Caps, Sorting, and the Regression-Discontinuity Design", *American Economic Review*, Vol. 99(1), pp. 179-215.
- [40] van der Klaauw, W. (2002) "Estimating the Effect of Financial Aid Offers on College Enrollment: A Regression-Discontinuity Approach", *International Economic Review*, Vol. 43(4), pp. 1249-1287.



## Appendix

Table A1: Descriptive statistics, sociodemographic characteristics by migration reason (in percentages)

	Economic migration	Family migration	Refugees
<i>Gender</i>			
Men	0.80	0.41	0.58
Women	0.20	0.59	0.42
<i>Age in 2010</i>			
18 to 29	0.22	0.46	0.46
30 to 39	0.50	0.37	0.34
40 to 49	0.21	0.12	0.12
50 or more	0.07	0.04	0.08
<i>Age at arrival</i>			
0 to 17	0.03	0.06	0.13
18 to 29	0.65	0.57	0.48
30 to 39	0.24	0.27	0.24
40 to 49	0.06	0.07	0.10
50 or more	0.01	0.02	0.05
<i>Duration of stay in France in 2010</i>			
Less than 2 years	0.09	0.60	0.30
From 2 to 4	0.10	0.12	0.51
From 5 to 9	0.59	0.22	0.16
10 years or more	0.22	0.07	0.04
<i>Nationality</i>			
Maghreb	0.20	0.46	0.01
Sub-Saharan Africa	0.51	0.21	0.23
Other areas in Africa	0.05	0.06	0.14
Asia	0.12	0.16	0.44
Europe (excluding France) and CIS	0.05	0.05	0.14
America and Oceania	0.07	0.06	0.03

Table A2: Descriptive statistics, labor market outcomes by gender (in percentages)

	All	Men	Women
<i>Labor force participation</i>			
Before migration	0.57	0.67	0.47
2010	0.69	0.88	0.51
2011	0.71	0.90	0.52
2013	0.77	0.94	0.62
<i>Employment rate</i>			
Before migration	0.52	0.63	0.43
2010	0.46	0.67	0.26
2011	0.54	0.74	0.35
2013	0.59	0.78	0.42
<i>Unemployment rate</i>			
Before migration	0.08	0.06	0.10
2010	0.33	0.24	0.48
2011	0.24	0.18	0.34
2013	0.21	0.15	0.29

Table A3: Positive Selection out of Training: Employment.

Positive Selection Out Of Language Training: Employment.								
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Local linear estimates			Parametric estimates				
	[35, 60]	[35, 60]	[30, 65]	[30, 65]				
<b>PART I: By education level.</b>								
Second stage regressions (dependent variable: employment dummy)								
Hours of French classes (/100)	-0.025 [0.128]	-0.153*** [0.052]	-0.013 [0.138]	-0.106* [0.056]	-0.008 [0.071]	0.054 [0.042]	-0.057 [0.123]	0.020 [0.079]
<b>PART II: By permanence in France.</b>								
Second stage regressions (dependent variable: employment dummy)								
Hours of French classes (/100)	-0.023 [0.129]	-0.137** [0.058]	-0.009 [0.140]	-0.062 [0.068]	-0.008 [0.071]	0.051 [0.043]	-0.053 [0.123]	0.013 [0.082]
<b>PART III: By language distance (Francophone vs. Non-Francophone).</b>								
Second stage regressions (dependent variable: employment dummy)								
Hours of French classes (/100)	-0.039 [0.123]	-0.141** [0.056]	-0.038 [0.124]	-0.079 [0.066]	-0.014 [0.069]	0.055 [0.041]	-0.082 [0.109]	0.017 [0.069]
<b>PART IV: By employment in 2010.</b>								
Second stage regressions (dependent variable: employment dummy)								
Hours of French classes (/100)	-0.024 [0.125]	-0.069 [0.057]	-0.007 [0.139]	-0.030 [0.061]	-0.006 [0.072]	0.051 [0.043]	-0.049 [0.123]	0.010 [0.084]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	51	51	58	58	1,734	1,734	1,734	1,734

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A4: Positive Selection out of Training: Full-time Contract.

Positive Selection Out Of Language Training: Full-time contract.								
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Local linear estimates			Parametric estimates				
	[35, 60]	[35, 60]	[30, 65]	[30, 65]				
<b>PART I: By education level.</b>								
Second stage regressions (dependent variable: full-time contract dummy)								
Hours of French classes (/100)	-0.187 [0.183]	0.337 [0.243]	-0.247 [0.188]	0.552 [0.369]	-0.349** [0.148]	-0.214*** [0.071]	-0.344 [0.217]	-0.349** [0.148]
<b>PART II: By permanence in France.</b>								
Second stage regressions (dependent variable: full-time contract dummy)								
Hours of French classes (/100)	-0.187 [0.183]	0.519 [0.492]	-0.242 [0.188]	0.918 [0.828]	-0.349** [0.149]	-0.209*** [0.070]	-0.340 [0.216]	-0.288** [0.131]
<b>PART III: By language distance (Francophone vs. Non-Francophone).</b>								
Second stage regressions (dependent variable: full-time contract dummy)								
Hours of French classes (/100)	-0.126 [0.205]	0.635 [0.657]	-0.176 [0.208]	1.177 [1.348]	-0.316** [0.144]	-0.201*** [0.069]	-0.255 [0.202]	-0.259** [0.129]
<b>PART IV: By employment in 2010.</b>								
Second stage regressions (dependent variable: full-time contract dummy)								
Hours of French classes (/100)	-0.178 [0.187]	0.370 [0.256]	-0.202 [0.189]	0.549 [0.361]	-0.340** [0.148]	-0.200*** [0.069]	-0.304 [0.210]	-0.252** [0.124]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	46	46	51	51	1,387	1,387	1,387	1,387

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A5: Positive Selection out of Training: Permanent Contract.

**Positive Selection Out Of Language Training: Permanent contract.**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Local linear estimates			Parametric estimates				
	[35, 60]	[35, 60]	[30, 65]	[30, 65]				
<b>PART I: By education level.</b>								
<b>Second stage regressions (dependent variable: permanent contract dummy)</b>								
Hours of French classes (/100)	-0.148 [0.190]	0.323 [0.225]	-0.117 [0.182]	0.220 [0.270]	-0.233* [0.122]	-0.074 [0.069]	-0.118 [0.142]	0.049 [0.124]
<b>PART II: By permanence in France.</b>								
<b>Second stage regressions (dependent variable: permanent contract dummy)</b>								
Hours of French classes (/100)	-0.140 [0.189]	0.343 [0.318]	-0.109 [0.182]	0.607 [0.558]	-0.231* [0.122]	-0.063 [0.067]	-0.110 [0.143]	0.092 [0.114]
<b>PART III: By language distance (Francophone vs. Non-Francophone).</b>								
<b>Second stage regressions (dependent variable: permanent contract dummy)</b>								
Hours of French classes (/100)	-0.162 [0.207]	0.263 [0.333]	-0.109 [0.191]	0.178 [0.430]	-0.228* [0.118]	-0.060 [0.066]	-0.103 [0.131]	0.087 [0.123]
<b>PART IV: By employment in 2010.</b>								
<b>Second stage regressions (dependent variable: permanent contract dummy)</b>								
Hours of French classes (/100)	-0.128 [0.181]	0.311 [0.216]	-0.114 [0.178]	0.154 [0.246]	-0.235* [0.124]	-0.069 [0.069]	-0.124 [0.146]	0.070 [0.124]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	46	46	51	51	1,387	1,387	1,387	1,387

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A6: Positive Selection out of Training: Informal Work.

**Positive Selection Out Of Language Training: Informal work.**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Local linear estimates			Parametric estimates				
	[35, 60]	[35, 60]	[30, 65]	[30, 65]				
<b>PART I: By education level.</b>								
<b>Second stage regressions (dependent variable: informal work dummy)</b>								
Hours of French classes (/100)	-0.119 [0.084]	0.046 [0.118]	-0.161* [0.085]	0.009 [0.147]	0.026 [0.066]	0.003 [0.031]	-0.111 [0.081]	-0.060 [0.059]
<b>PART II: By permanence in France.</b>								
<b>Second stage regressions (dependent variable: informal work dummy)</b>								
Hours of French classes (/100)	-0.122 [0.084]	-0.077 [0.199]	-0.163* [0.086]	-0.235 [0.281]	0.026 [0.066]	-0.002 [0.031]	-0.113 [0.084]	-0.079 [0.059]
<b>PART III: By language distance (Francophone vs. Non-Francophone).</b>								
<b>Second stage regressions (dependent variable: informal work dummy)</b>								
Hours of French classes (/100)	-0.094 [0.090]	0.102 [0.221]	-0.146* [0.080]	0.057 [0.242]	0.029 [0.065]	0.005 [0.031]	-0.093 [0.076]	-0.054 [0.057]
<b>PART IV: By employment in 2010.</b>								
<b>Second stage regressions (dependent variable: informal work dummy)</b>								
Hours of French classes (/100)	-0.124* [0.075]	0.029 [0.107]	-0.142* [0.076]	-0.004 [0.134]	0.032 [0.067]	0.004 [0.032]	-0.091 [0.079]	-0.057 [0.059]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	46	46	51	51	1,387	1,387	1,387	1,387

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A7: Positive Selection out of Training: Income per Household Individual.

Positive Selection Out Of Language Training: Income per household individual.								
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Local linear estimates					Parametric estimates		
	[35, 60]	[35, 60]	[30, 65]	[30, 65]				
<b>PART I: By education level.</b>								
<b>Second stage regressions (dependent variable: income per household individual)</b>								
Hours of French classes (/100)	31.891 [101.931]	186.510** [89.172]	-47.931 [93.283]	129.194 [81.493]	-59.113 [90.052]	-107.998** [53.815]	-154.781 [100.749]	-52.557 [62.384]
<b>PART II: By permanence in France.</b>								
<b>Second stage regressions (dependent variable: income per household individual)</b>								
Hours of French classes (/100)	28.434 [102.966]	185.638** [91.856]	-49.246 [96.112]	119.001 [83.481]	-59.864 [91.030]	-112.460** [53.180]	-157.779 [106.443]	-63.267 [63.685]
<b>PART III: By language distance (Francophone vs. Non-Francophone).</b>								
<b>Second stage regressions (dependent variable: income per household individual)</b>								
Hours of French classes (/100)	34.370 [103.709]	160.564* [83.648]	-73.831 [92.620]	82.780 [76.688]	-68.049 [87.299]	-110.177** [53.040]	-170.971* [93.854]	-64.454 [57.369]
<b>PART VI: By employment in 2010.</b>								
<b>Second stage regressions (dependent variable: income per household individual)</b>								
Hours of French classes (/100)	2.619 [94.010]	122.654* [73.619]	-41.823 [89.838]	104.798 [75.421]	-54.285 [91.370]	-115.815** [53.944]	-145.177 [100.560]	-69.529 [62.952]
Polynomial of degree	1	1	1	1	1	1	2	2
Control variables	No	Yes	No	Yes	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes	No	Yes	No	Yes
Observations	61	61	69	69	2,040	2,040	2,040	2,040

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A8: Assigned French Language Training and Other Measures of Labor Market Integration. Excluding Test Result = 0.

Assigned French Language Training and other measures of Labor Market Integration. Excluding Test Result = 0.				
	(1)	(2)	(3)	(4)
		Parametric estimates		
<b>Panel A: Dependent variable: employment (conditional on labor force participation). Second stage regressions.</b>				
Hours of French classes (/100)	-0.058 [0.065]	0.031 [0.039]	-0.052 [0.103]	0.036 [0.069]
Observations	1,690	1,690	1,690	1,690
<b>Panel B: Dependent variable: full-time work (conditional on employment). Second stage regressions.</b>				
Hours of French classes (/100)	-0.348** [0.157]	-0.200*** [0.066]	-0.212 [0.190]	-0.235* [0.128]
Observations	1,357	1,357	1,357	1,357
<b>Panel C: Dependent variable: permanent contract (conditional on employment). Second stage regressions.</b>				
Hours of French classes (/100)	-0.239* [0.124]	-0.076 [0.073]	-0.037 [0.127]	0.181 [0.123]
Observations	1,357	1,357	1,357	1,357
<b>Panel D: Dependent variable: informal work (conditional on employment). Second stage regressions.</b>				
Hours of French classes (/100)	0.029 [0.058]	0.006 [0.028]	-0.142 [0.091]	-0.134** [0.064]
Observations	1,357	1,357	1,357	1,357
<b>Panel E: Dependent variable: income per household individual. Second stage regressions.</b>				
Hours of French classes (/100)	-73.174 [80.126]	-120.737** [49.976]	-151.557 [93.462]	-19.797 [58.151]
Observations	1,967	1,967	1,967	1,967
Polynomial of degree	1	1	2	2
Control variables	No	Yes	No	Yes
Country of origin fixed effects	No	Yes	No	Yes

Robust standard errors in brackets, clustered by country of origin and test scores of the initial language exam.  
 The columns include functions in the test scores of the initial language exam (distance from the cutoff value) of degree 1 and 2, interacted with I(test score initial language exam < 50).  
 The set of control variable includes: education level in years, age, age squared, a dummy variable equal to 1 if the immigrant is resident in the region "Ile-de-France" (the region around Paris), a dummy variable equal to 1 if the immigrant is married, a dummy variable equal to 1 if the immigrant is male, the number of children in the household, the total number of individuals in the household, the number of years that the migrant has already spent in France, dummy variables indicating the reason to migrate (labor migration, refugee, other. Family migration is the reference category.), a dummy variable equal to one if the individual was employed in 2010 and country of origin fixed effects.  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1