

The Cultural Assimilation of Individualism and Preferences for Redistribution*

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Abstract

I analyze the relationship between individualism and preferences for redistribution, using variation in immigrants' countries of origin to capture the impact of cultural values and beliefs on personal attitudes towards income redistribution and equality. Using global individual-level survey data for more than one million individuals (including 65,000 migrants) in a large number of countries around the world, I find strong support for the hypothesis that more individualistic cultures are associated with lower preferences for redistribution. At the same time, cultural assimilation in this dimension seems to take place relatively fast, where the impact of the destination culture starts to dominate the origin culture when an individual has lived as long in the country of destination as she did in her country of origin. Moreover, I find no statistically significant effect of the origin culture on an individual's preferences for redistribution if migration took place before the age of 10. The results are confirmed using a variety of robustness checks, including the grammatical rule of a pronoun drop as an instrumental variable.

Keywords: Individualism, Preferences for redistribution, Cultural assimilation, Migration

JEL Classification Codes: H23, O15, Z13

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1 Introduction

There is large variation in the amount of income equality and redistribution people support around the world, both on the individual and country level (Alesina & Angeletos, 2005; Almås, Cappelen, Sørensen & Tungodden, 2020). Recent studies have found culture¹ to be an important determinant of such variations in preferences for redistribution (Luttmer & Singhal, 2011; Alesina & Giuliano, 2015). Less studied, however, is which cultural dimensions that affects our preferences and attitudes towards equality and redistribution? In this study, I will focus on the relationship between one particular cultural dimension, individualism versus collectivism, and preferences for redistribution.

In a recent paper by Bazzi, Fiszbein & Gebresilasse (2020), they find that the expansion of the U.S. border fostered "rugged individualism", i.e., a combination of individualism and opposition to government intervention and redistribution. In this paper, I will empirically test whether this combination of individualism and opposition to redistribution is a more general phenomenon that also holds among a global sample of migrants around the world today. To my knowledge, this study is the first to test this relationship on a broader scale.

A second, related question is what happens with people's preferences after they have migrated into a new culture. Will their preferences and attitudes converge with the new cultural environment and, if so, how fast? By analyzing people that migrated at different ages and have lived differently long in their country of destination, I will study the cultural assimilation of individualism and preferences for redistribution by comparing the cultural origin versus cultural destination impact over the life course of migrants.

While individualism versus collectivism has been considered a main dimension of cross-country cultural variation in the psychology literature since long (see, e.g., Hofstede, Hofstede & Minkov, 2010; Heine, 2020), and although Greif (1994) argued for its economic importance more than 25 years ago, it is only more recently that this cultural dimension

¹In this paper, I will follow Fernández (2011) and Alesina & Giuliano (2015) use of the term culture when referring to, e.g., social values and beliefs. This concept, however, is also closely related to what, e.g., North (1991) and Williamson (2000) refer to as informal institutions. Culture can be defined as "the set of values and beliefs people have about how the world (both nature and society) works as well as the norms of behavior derived from that set of values" (Gorodnichenko & Roland, 2017).

has gained recognition in empirical economics (e.g., [Gorodnichenko & Roland, 2011, 2017, 2020](#)). Also, most previous studies on individualism highlight its (positive) correlation with, for instance, economic growth (e.g., [Gorodnichenko & Roland, 2011](#)). At the same time, however, whether a society is more individualistic or collectivistic, i.e., whether people's self-image is defined in terms of "I" or "we", is also likely to affect how individuals value equality. This potential link, I believe, is something that has not been investigated before, and studying this relationship is thus important for our understanding of the cultural roots of preference differences. Moreover, these preferences will eventually also have implications for actual redistribution and welfare around the world.

In order to test whether individuals from more individualistic cultures prefer less income redistribution,² I will use variation in immigrants' country of origin to capture the impact of culture on individual preferences (i.e., using the so-called "epidemiological" approach), using individual-level survey data from a large set of countries around the world using the integrated *World Values Survey* ([WVS, 2016](#)) and *European Values Study* ([EVS, 2016](#)), and the *European Social Survey* ([ESS, 2016](#)). As robustness checks, I will also apply matching estimators as well as an instrumental variables (IV) approach where I use the pronoun drop dummy from [Abdurazokzoda & Davis \(2016\)](#)'s new linguistic dataset as an instrument for collectivism. Doing so, I find a robust and statistically significant negative relationship between individualism and preferences for redistribution. Heterogeneity analyses confirm this association for both individuals born in another country and with another citizenship, while there seems to be assimilation into the new cultural environment over time as the impact is not persistent for second-generation immigrants. When individuals have spent approximately half their life in the new country, the impact of the country-of-residence culture starts to dominate that of their country of origin. Moreover, I find no statistically significant impact of the culture of origin if migration took place before the age of 10.

The rest of the paper is organized as follows. In the next section I will give a brief

²While there is a difference between preferences for redistribution and preferences for equality, both conceptually and empirically, in my main analysis I will use these concepts interchangeably as the survey data that I use include measures of both. In the sensitivity analysis, however, I will make a distinction between these two questions.

overview of some previous empirical research that has been conducted on this topic. Thereafter, I will present the empirical approach and data used in this paper, followed by its main results, including an analysis of cultural assimilation. Finally, I will also present some robustness and heterogeneity analyses, before I conclude.

2 Previous Research and Hypothesis

Quite a few earlier studies have tried to use cross-country data to analyze the determinants of preferences for redistribution or equality. However, a problem with many of these studies is that they use too aggregated data and thereby risk averaging away potentially important individual determinants of preferences. Also, the relationships could be different on the cross-country and individual levels. As a response to this, more recent studies have instead used individual-level survey data and, as such, been able to also take individual characteristics into account. Yet, when trying to establish causal relationships a number of problems, including endogeneity issues such as reverse causality, simultaneity and omitted variables, remain. As a potential solution to some of these problems, the so-called epidemiological approach has become popular over the last years. In such a study, [Luttmer & Singhal \(2011\)](#) find that immigrants' redistributive preferences are positively related to the average preference in their birth countries, and that cultural determinants of preferences for redistribution are persistent across generations. This means that redistributive preferences cannot be fully explained by economic self-interest or by the current economic, political or social environment. In other words, culture seems to matter. But which aspects of culture are important in shaping preferences for redistribution and equality? This is an important question that remains yet to be answered.

The individualism-collectivism cleavage is one such particular aspect of culture, which has been claimed to be the most important cross-country dimension in cultural psychology ([Heine, 2020](#)), as well as the primary cultural dimension affecting long-run economic growth ([Gorodnichenko & Roland, 2011, 2012, 2017](#)). While individualism emphasizes personal freedom and achievement, collectivism emphasizes embeddedness of individu-

als in larger groups. Because an individualistic culture implies stronger preferences for personal freedom, it seems plausible that individuals in such a culture should prefer less income redistribution, and possibly also less income equality. A collectivistic culture, on the other hand, is associated with considerations beyond the individual self, i.e., for the group, and is thus likely to imply more egalitarian preferences. My hypothesis is thus that more individualistic societies should foster preferences for more income inequality, and vice versa.

Other potential determinants of redistributive and equality preferences that have been found significantly (and negatively) associated with preferences for redistribution in previous studies include the individual characteristics of age, being male, income, right-wing ideology, education and employment (Alesina & Giuliano, 2011). Moreover, preferences for redistribution have also been found to be affected by country-level and time-specific determinants such as political regimes (Alesina & Fuchs-Schündeln, 2007), macroeconomic shocks (Giuliano & Spilimbergo, 2014) and changes in income inequality (Olivera, 2015; Schmidt-Catran, 2016). Finally, social trust has been found to affect income equality in previous studies (e.g., Bergh & Bjørnskov, 2014), and a potential mechanism for this relationship could be via redistributive preferences.

3 Data and Empirical Approach

3.1 Epidemiological approach

The idea behind the epidemiological approach is that culture affects prior beliefs, which in turn affect economic outcomes (Guiso, Sapienza & Zingales, 2006). As such, it analyzes the variation in outcomes across different (first- or second-generation) immigrant groups residing in the same country, thus making it possible to separate the impact of culture from the, otherwise endogenously determined, economic and institutional environment. The assumptions underlying this approach is that cultural beliefs vary across immigrant groups in a systematic fashion reflecting culture in the country of origin, and that individuals who live in the same country face similar economic and formal institu-

tional environments (Fernández, 2011). As an example of the epidemiological approach, Alesina & Giuliano (2011) and Luttmer & Singhal (2011) have found that culture, as measured by the mean preferences for redistribution in the immigrants' country of origin, appears to be an important determinant of preferences for redistribution.

Following this approach, my baseline estimation equation is given by:

$$Preferences_{ijct} = \beta_0 + \beta_1 IDV_c + \beta_2 Z_c + \beta_3 X_i + \gamma_j + \mu_t + \epsilon_{ijct},$$

where $Preferences_{ijct}$ is the preferences for redistribution of individual i , living in country j and coming from country c ($c \neq j$); in year t ; IDV_c is the individualism index in the individual's country of origin; Z_c is a vector of country-of-origin-level controls; X_i is a vector of individual controls; γ_j and μ_t are country of residence and year fixed effects, respectively; and ϵ_{ijct} is an error term. The country of residence fixed effect captures the institutional environment and all other unobserved characteristics that apply to all individuals living in that country. It also implies that the cultural variable captures the difference between the social beliefs in the individual's country of origin relative to the country of residence (i.e., the cultural component; see, e.g., Dinesen, 2012). The regressions are run using ordinary least squares (OLS), but using ordered logistic or probit regressions yields qualitatively the same results. The results are shown with robust standard errors clustered on country of origin, but the results also hold with country-of-residence clustering.

In a robustness analysis, I will also use propensity score matching to compare individuals with an individualistic culture to similar individuals with the main difference being that they have a collectivistic culture instead. This approach has previously been used in a similar context by, e.g., Dinesen (2012) comparing the level of trust of migrants and comparable non-migrants. Moreover, I will also use an IV approach as an alternative identification strategy trying to disentangle the effect of individualism from other cultural components. As instrument for individualism-collectivism I will then use a linguistic measure of the grammatical rule on pronoun drop, which was first collected by Kashima & Kashima (1998) and recently expanded by Abdurazokzoda & Davis (2016). As an example of such a pronoun drop you can, e.g., in Spanish say both "yo hablo"

(*"I speak"*) or only *"hablo"* (dropping the subject pronoun *"yo"*), while such a pronoun drop is not permitted in, e.g., English. The intuition behind this instrument is that more individualistic societies tend to emphasize the importance of the individual in the context of speech and thus have kept the pronoun, while more collectivistic societies more often have dropped it. Previous studies using the pronoun drop as similar instrument are [Licht, Goldschmidt & Schwartz \(2007\)](#), [Tabellini \(2008\)](#) and [Alesina & Giuliano \(2010\)](#). This linguistic feature is then assumed to affect preferences for redistribution only through its relationship with individualism.

3.2 Global survey data

Most previous studies analyzing the determinants of preferences for redistribution or equality only use data from one specific survey, country or region. I broaden this approach by using a combined dataset of the integrated *World Values Survey* ([WVS, 2016](#)) and *European Values Study* ([EVS, 2016](#)) and the *European Social Survey* ([ESS, 2016](#)), thus obtaining a wide set of countries and individuals from all around the world. The coverage of this dataset is shown in [Table 1](#). In the full sample, the [WVS \(2016\)](#) includes 341,271 individuals in 98 countries over the years 1981–2014, the [EVS \(2016\)](#) includes 164,997 individuals in 46 countries for 1981–2009, and the [ESS \(2016\)](#) includes 336,964 individuals in 36 countries for 2002–2014.

Most of my analysis will focus on the sample of all individuals with another cultural origin than their country of residence, in which I will include both first-generation immigrants (i.e., individuals with another nationality or country of birth) and second-generation immigrants (i.e., individuals whose mother and/or father has another country of origin). In some heterogeneity analyses, however, I will also compare and analyze these different samples and datasets separately.

As dependent variable, I will use individuals' responses to the survey question on income equality values, i.e., self-selection on a 10-point scale ranging from *"We need larger income differences as incentives"* to *"Incomes should be made more equal"* ([EVS, 2016](#); [WVS, 2016](#)). In the [ESS \(2016\)](#), this question is phrased slightly differently, namely as

Table 1: Coverage of the dataset.

| | Individuals | Residence countries | Origin countries | Years |
|---|-------------|---------------------|------------------|-----------|
| Migrant sample: other cultural origin | 65,426 | 78 | 188 | 2002–2020 |
| <i>First-generation immigrants</i> | | | | |
| Other citizenship | 18,645 | 45 | 168 | 2002–2018 |
| Other country of birth | 49,625 | 76 | 187 | 2002–2020 |
| <i>Second-generation immigrants</i> | | | | |
| Other origin both parents | 51,294 | 75 | 182 | 2004–2020 |
| Other origin mother | 65,352 | 76 | 184 | 2004–2020 |
| Other origin father | 66,878 | 80 | 183 | 2004–2020 |
| Full sample: both migrants and non-migrants | 1,079,107 | 113 | 188 | 1981–2020 |

Note: Other cultural origin is defined as having another citizenship, another country of birth or both parents having another country of birth than the country of residence. One individual can belong to multiple groups.

Sources: ESS (2016); EVS (2016); WVS (2016).

"The government should take measures to reduce differences in income levels", with selection on a 5-point scale ranging from "Agree strongly" to "Disagree strongly". Conceptually, the EVS/WVS question is thus closer to the concept of income equality preferences, while the ESS question is closer to that of preferences for income redistribution. In my baseline analysis, I will use both sources and thus recode this variable into an index ranging from 0 to 100, where a higher value indicates stronger preferences for income equality or redistribution, and vice versa. Sensitivity analyses, however, show that the results do not depend on the wording of this question and hold for each survey separately. Country coverage and the average preferences for redistribution are illustrated in Figure A1 in the Appendix.

The individualism-collectivism explanatory variable is collected from Hofstede et al. (2010) and their later extensions,³ whose individualism index is the most commonly used empirical measure of this cultural dimension (Alesina & Giuliano, 2015; Gorodnichenko & Roland, 2017). This dimension has also been included in recent research using the epidemiological approach, albeit looking at other outcomes (Berggren, Ljunge & Nilsson, 2019; Ljunge, 2017). The individualism index is given at the country level for 102 countries (see Figure 1 for country coverage and individualism values) and assumed to be

³Available at <https://geerthofstede.com> (last accessed: September 12, 2016)

constant over the analyzed time period, which should be reasonable given that cultures usually change only slowly over time (Williamson, 2000). The index is based on factor analysis using survey questions (initially for IBM employees, but later expanded) and has been validated in a number of studies (see, e.g., Gorodnichenko & Roland, 2017; Hofstede et al., 2010).⁴ It ranges from 0 to 100, with 0 representing maximum collectivism, i.e., "a society in which people from birth onwards are integrated into strong, cohesive in-groups, which continue to protect them throughout their lifetime in exchange for unquestioning loyalty", and 100 maximum individualism, i.e., "a society in which the ties between individuals are loose: a person is expected to look after himself or herself and his or her immediate family only" (Hofstede & Minkov, 2013). In the main analysis, immigrants are assigned the individualism index value of their country of origin (i.e., country of nationality, country of birth, mother's or father's country of origin, if different than country of residence).⁵ In the heterogeneity analysis, however, I analyze these different samples separately.

The individual-level control variables taken from the WVS (2016), the EVS (2016) and the ESS (2016) include the (recoded) survey measures of trust (1 meaning that the individual answered "Most people can be trusted", in contrast to 0 "Can't be too careful"), satisfaction with life (ranging from 0 "Dissatisfied" to 100 "Satisfied"), self positioning in political scale (ranging from 0 "Left" to 100 "Right"), highest educational level attained (ranging from 0 "Inadequately completed elementary education" to 100 "University with degree"), employment status (where 0 means "Unemployed", 1 "Other"

⁴The index formula used by Hofstede et al. (2010) to calculate the individualism index (IDV) is given by:

$$IDV = 35(\text{Mean}Q4 - \text{Mean}Q1) + 35(\text{Mean}Q9 - \text{Mean}Q6) + \text{Constant},$$

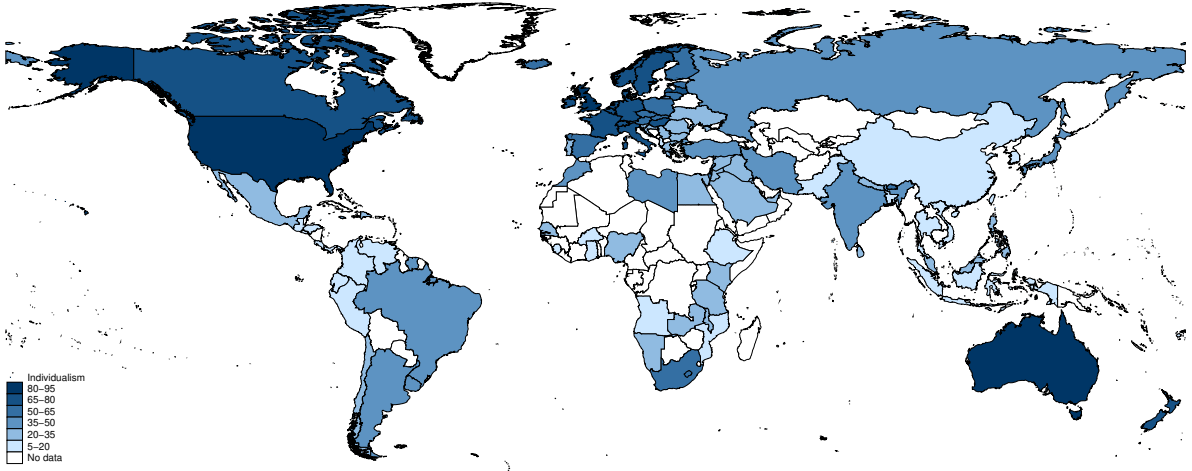
where $\text{Mean}QX$ is the mean score of question X in the following:

- "In choosing an ideal job, how important would it be for you to:*
- 1) have sufficient time for your personal or home life;*
 - 4) have security of employment;*
 - 6) do work that is interesting;*
 - 9) have a job respected by your family and friends",*

ranked on a 5-point scale, ranging from 1 "of utmost importance" to 5 "of very little or no importance".

⁵If an individual has both another nationality and country of birth, I simply use their average. Similarly, for second-generation immigrants, I use the average value of both parents. If an individual is both a first- and second-generation immigrant, I use the value of his or her own country of origin.

Figure 1: Individualism versus collectivism around the world.



Source: Hofstede et al. (2010).

and 2 "*Employed*"), monthly household income (in constant euros), a sex dummy (where 0 indicates male and 1 indicates female), age, and number of years lived in country (grouped into less than 1 year, 1–5 years, 6–10 years, 11–20 years, and more than 20 years).

In the baseline specification, I will control for the level of inequality and income in the country of origin and use fixed-effects for the country of residence. In alternative specifications, however, I will control for a broader set of variables in the country of origin, including the mean level of social trust and equality preferences, ethnolinguistic fractionalization and democratic rights. The sources for these country-level control variables are the following: the actual level of income inequality is measured by the Gini coefficient as collected by Milanovic (2016)'s *All the Ginis* (ALG) dataset; the country-level income is measured by the log of GDP per capita (in PPP-adjusted constant 2011 international dollars) taken from the World Bank (2016)'s *World Development Indicators* (WDI); the average ethnic, linguistic and religious fractionalization is measured by Alesina, Devleeschauwer, Easterly, Kurlat & Wacziarg (2003); and, as an indicator of democracy and autocracy, I use the revised combined polity score (rescaled into a 0–100 index) from the Polity IV Project (Marshall, Gurr & Jaggers, 2016). Moreover, as further controls, the survey values for social trust and equality preferences above are averaged at the country level. The pronoun drop instrument is taken from Abdurazokzoda & Davis

(2016) new linguistic dataset (where 1 indicates that the language allows pronoun drop and 0 that it does not), covering 56 languages in 94 countries (the country-level averages of this dummy is illustrated in Figure A2 in the Appendix). Some summary statistics for the different variables and samples are presented in Table 2.

From Table 2, the immigrant sample seems fairly representative for the full sample, even though it covers a smaller sample of residence countries and years. Pairwise correlations for the individual and country-of-origin-level characteristics are presented in Table A1 in the Appendix. These correlations indicate that, on the individual level, preferences for more equal income distributions seem to be correlated with higher trust, lower life satisfaction, more left-wing political preferences, lower educational, employment and income levels, time spent in the new country, and being female and older. On the country level, focusing on the cultural component, individual preferences for redistribution and equality seem to be negatively related to individualism, social trust, actual income equality, GDP per capita, fractionalization and democratic rights in the immigrants' country of origin. The correlation between individualism index in country of origin and individual preferences for redistribution is illustrated in the binned scatterplot in Figure 2 (corresponding correlations separated into the different immigrant samples are shown in Figure A3 in the Appendix). However, these are only simple correlations and, hence, I turn now to the regression results.

4 Main results

4.1 Relation between individualism and redistributive preferences

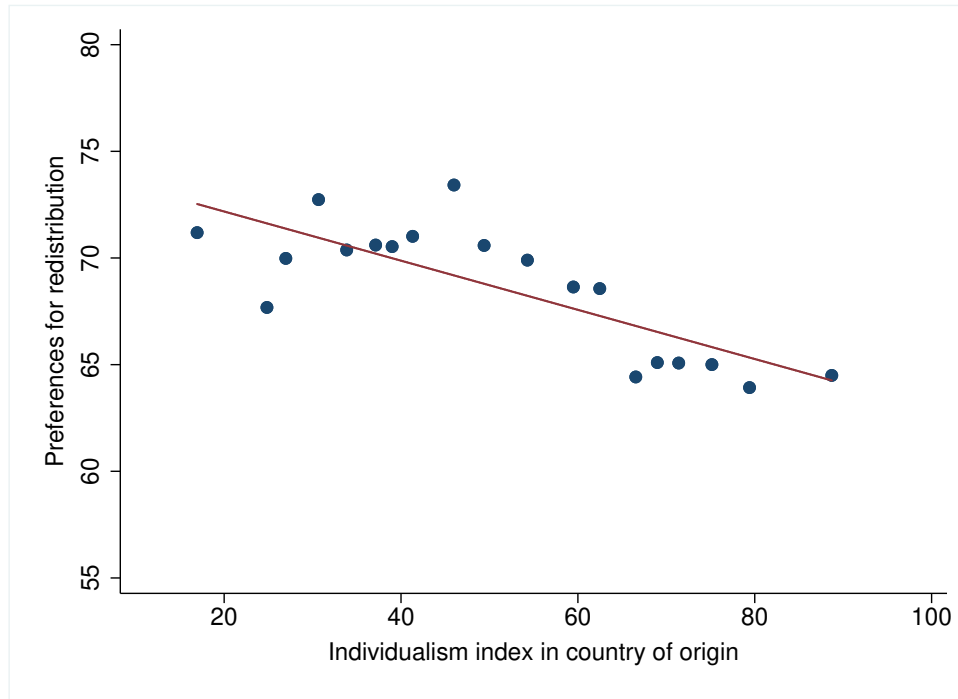
I first run a standard OLS regression estimating individuals' preferences for redistribution by the level of individualism in their country of residence (including the full sample of both immigrants and non-immigrants) and a set of individual and country-level controls. These results are presented in Table A2 in the Appendix (together with the same estimation using ordered logistic and ordered probit regressions, respectively, instead of OLS), and show a negative and statistically significant relationship between individualism and

Table 2: Summary statistics.

| | Obs. | Mean | Std. Dev. | Min | Max |
|---|-----------|--------|-----------|-------|---------|
| <i>Migrant sample</i> | | | | | |
| Incomes should be made more equal | 15,806 | 48.0 | 31.8 | 0 | 100 |
| Government should reduce income differences | 47,712 | 72.7 | 25.7 | 0 | 100 |
| Individualism index (origin country) | 62,344 | 44.7 | 21.0 | 6 | 91 |
| Individualism index (residence country) | 62,265 | 59.4 | 17.5 | 6 | 91 |
| Gini coefficient (origin country) | 64,541 | 35.9 | 6.8 | 19 | 77 |
| Gini coefficient (residence country) | 63,689 | 33.0 | 5.0 | 23 | 56 |
| GDP per capita (origin country) | 64,143 | 25,027 | 15,493 | 715 | 131,908 |
| GDP per capita (residence country) | 64,467 | 44,480 | 20,185 | 2,836 | 123,965 |
| Polity score (origin country) | 65,120 | 5.25 | 5.68 | -10 | 10 |
| Polity score (residence country) | 62,604 | 8.77 | 2.19 | -7 | 10 |
| Fractionalization (origin country) | 65,118 | 0.344 | 0.169 | 0.005 | 0.840 |
| Fractionalization (residence country) | 65,074 | 0.355 | 0.149 | 0.021 | 0.814 |
| Most people can be trusted | 64,705 | 46.1 | 32.3 | 0 | 100 |
| Left–right political scale | 53,939 | 49.1 | 23.0 | 0 | 100 |
| Household income decile | 45,162 | 45.3 | 29.1 | 0 | 100 |
| Highest educational level | 59,751 | 2.23 | 0.62 | 1 | 3 |
| Sex (1 Female, 2 Male) | 65,385 | 1.45 | 0.50 | 1 | 2 |
| Age | 64,987 | 47.0 | 17.8 | 12 | 114 |
| Age at migration | 41,818 | 21.0 | 12.7 | 0 | 110 |
| Years since migration | 42,025 | 26.8 | 18.2 | 0 | 92 |
| Year | 65,426 | 2011 | 5 | 2002 | 2020 |
| Pronoun drop dummy | 51,482 | 0.30 | 0.46 | 0 | 1 |
| <i>Full sample</i> | | | | | |
| Incomes should be made more equal | 569,928 | 47.0 | 33.3 | 0 | 100 |
| Government should reduce income differences | 429,765 | 72.6 | 25.5 | 0 | 100 |
| Individualism index (residence country) | 1,041,315 | 50.9 | 22.0 | 6 | 91 |
| Gini coefficient (residence country) | 1,063,119 | 34.9 | 8.2 | 20 | 66 |
| GDP per capita (residence country) | 1,069,232 | 30,353 | 18,477 | 1,008 | 123,965 |
| Polity score (residence country) | 1,057,137 | 7.42 | 4.50 | -10 | 10 |
| Fractionalization (residence country) | 1,076,211 | 0.342 | 0.173 | 0.005 | 0.829 |
| Most people can be trusted | 1,048,229 | 37.2 | 39.4 | 0 | 100 |
| Left–right political scale | 842,699 | 51.2 | 24.3 | 0 | 100 |
| Household income decile | 841,831 | 42.1 | 27.2 | 0 | 100 |
| Highest educational level | 900,555 | 2.10 | 0.64 | 1 | 3 |
| Sex (1 Female, 2 Male) | 1,073,873 | 1.47 | 0.50 | 1 | 2 |
| Age | 1,071,857 | 44.9 | 17.8 | 12 | 123 |
| Year | 1,079,107 | 2007 | 9 | 1981 | 2020 |
| Pronoun drop dummy | 513,796 | 0.51 | 0.50 | 0 | 1 |

Sources: Alesina et al. (2003); ESS (2016); EVS (2016); Hofstede et al. (2010); Marshall et al. (2016); Milanovic (2016); World Bank (2016); WVS (2016).

Figure 2: Correlation between cultural individualism and individual preferences for redistribution: binned scatterplot.



Note: Bins based on 55,334 observations in total.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); WVS (2016).

preferences for redistribution also when controlling for a large number of individual and country-of-residence-level characteristics. Most of the individual variables also remain statistically significant and in the same direction as the pairwise correlations, except for trust, which is now found to have a negative conditional correlation with redistributive preferences. When controlling for the other variables, income inequality in the country of residence is positively related to preferences for redistribution and equality, while country-of-residence-level income has a positive, but not statistically significant, association with preferences for income redistribution.

However, since all individuals living in a country are assumed to have the same individualism-collectivism cultural beliefs in the regression above, I cannot control for country fixed effects, and this specification could potentially suffer from a number of endogeneity issues. Thus, I now turn to the epidemiological approach, exploiting variation in immigrants' country of origin to better capture variation in the cultural dimension of individualism versus collectivism within the country of residence, i.e., applying country-

of-residence fixed effects to control for the economic and institutional environment in which these individuals live. In particular, this should solve for any unobservable characteristics at the country-of-residence level, as well as for potential reverse causality since the individual preferences of a person living in a new country are not very likely to affect the individualism-collectivism ranking of his or her country of origin. These OLS regression results are presented in Table 3.⁶

The results in Table 3 show a negative and statistically significant relationship between individualistic cultural beliefs in the country of origin and immigrants' preferences for income redistribution, even after controlling for their country of residence. That is, more individualistic cultures seem to be associated with less egalitarian values, where an increase of 10 percentage points on the individualism index is associated with a decrease of approximately 0.6–0.8 percentage points on the preferences for redistribution scale. Also, preferences for redistribution are found to be statistically significantly related to having lower life satisfaction, political preferences more to the left, lower levels of education, employment and income, more years spent in the new country, and being female and old. It could also be noted that none of the other country-of-origin-level variables are statistically significant (except for fractionalization, which is marginally significant), when controlling for the cultural individualism impact. The last column of Table 3 shows the same results but where the values of all variables have been standardized to having mean 0 and standard deviation 1. In other words, a one standard deviation increase of the individualism index (which corresponds, e.g., to the difference between Sweden with an individualism index value of 71 and the United States with a value of 91) is associated with a 0.05 standard deviation decrease of the preferences for redistribution measure (corresponding roughly to, e.g., the difference between the average redistributive preferences value in Sweden, which is 62.3, and in Poland, which is 60.6). This standardized coefficient magnitude is similar to that of household income and the life satisfaction measure, and the only variables with larger standardized coefficients are political preferences and the education level.

⁶Running ordered logit or probit regressions instead of OLS yields qualitatively the same results (results available upon request).

Table 3: Baseline OLS regression results, immigrant sample.

| Preferences for redistribution | (1) | (2) | (3) | <i>Std. Coef.</i> |
|---------------------------------------|----------------------|----------------------|----------------------|----------------------|
| <i>Origin country</i> | | | | |
| Individualism index | -0.057*** (0.018) | -0.068*** (0.025) | -0.081*** (0.027) | -0.049*** (0.017) |
| Gini coefficient | | -0.074 (0.057) | -0.015 (0.079) | -0.003 (0.015) |
| Log GDP per capita | | 0.672 (0.568) | -0.597 (0.879) | -0.014 (0.020) |
| Mean social trust | | | 0.052 (0.035) | 0.023 (0.015) |
| Mean preferences for redistribution | | | 0.015 (0.049) | 0.004 (0.014) |
| Fractionalization | | | -0.064* (0.038) | -0.030* (0.018) |
| Polity score | | | 0.030 (0.023) | 0.024 (0.019) |
| <i>Individual characteristics</i> | | | | |
| Trust value | | 1.219 (0.838) | 1.133 (0.862) | 0.017 (0.013) |
| Life satisfaction value | | -0.054*** (0.014) | -0.060*** (0.014) | -0.047*** (0.011) |
| Political left–right scale | | -0.175*** (0.019) | -0.174*** (0.019) | -0.128*** (0.014) |
| Education level | | -0.080*** (0.013) | -0.078*** (0.013) | -0.074*** (0.012) |
| Employment status | | -0.674* (0.384) | -0.809** (0.388) | -0.016** (0.007) |
| Household income | | -0.001*** (0.000) | -0.001*** (0.000) | -0.052*** (0.014) |
| Time in new country | | 1.243*** (0.331) | 1.184*** (0.337) | 0.039*** (0.011) |
| Sex | 2.584*** (0.207) | 1.847*** (0.472) | 1.929*** (0.480) | 0.029*** (0.007) |
| Age | 0.107*** (0.014) | 0.050*** (0.018) | 0.048*** (0.018) | 0.029*** (0.010) |
| Residence country FE | Yes | Yes | Yes | Yes |
| Year FE | Yes | Yes | Yes | Yes |
| Number of obs. | 55,085 | 12,143 | 11,805 | 11,805 |
| R-squared | 0.117 | 0.153 | 0.158 | 0.158 |

Notes: Robust standard errors clustered on origin country in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: Alesina et al. (2003); ESS (2016); EVS (2016); Hofstede et al. (2010); Marshall et al. (2016); Milanovic (2016); World Bank (2016); WVS (2016).

4.2 Cultural assimilation analysis

Separating the sample into first- and second-generation immigrants (see Section 5.2), the association between country-of-origin individualism and individual preferences for income inequality is not found statistically significant among second-generation immigrants. The corresponding results for an even finer division of the immigrant subsamples are shown in Table A3 in the Appendix, and the standardized estimated coefficients of these regressions are illustrated in Figure 3. A potential explanation for this is that, with time spent in the new cultural environment, immigrants might also adapt the culture of their new country of residence (i.e., direct horizontal socialization, as opposed to vertical parental socialization; see, e.g., Bisin & Verdier, 2011). In order to further check whether or not this seems to be the case, I will, instead of the country-of-residence fixed effects, include the country-of-residence individualism index as well as income and inequality controls in my baseline regression. Because such cultural assimilation is also likely to depend on the relative time that the individual has lived in the new environment, I will calculate the country-of-residence life-share as the total number of years lived in the country-of-residence divided by the individual's age. Since the number of observations for each year lived in the country is quite limited, I will group this variable into three categories: those immigrants that have lived i) less than one third, ii) between one and two thirds, and iii) more than two thirds of the life in the new country of residence. The results for these three groups are presented in Table 4 and illustrated by their standardized coefficient estimates in Figure 4.

While these cultural assimilation results are admittedly crude, I believe that they do give some support to the idea that immigrants are assimilated or adapted to their new cultural environment quite "rapidly", i.e., after spending approximately the same amount of time in the new country as spent in the origin country. Moreover, the country of residence impact seems to be relatively stronger, with a coefficient size that is almost the double, compared to the culture of origin impact.

Finally, I will also perform the same cultural assimilation analysis as above, but instead separating the sample by the individuals' age when they migrated. These results

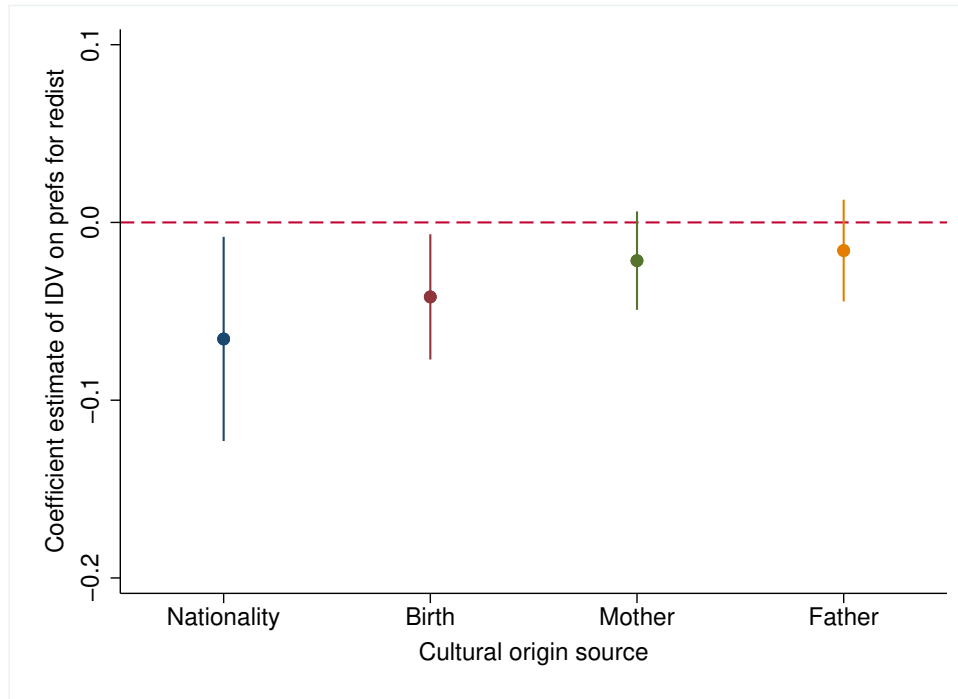
Table 4: Comparing country-of-origin versus country-of-residence cultural impact.

| Preferences for redistribution | <i>Life share in residence country</i> | | | | | |
|-----------------------------------|--|----------------------|----------------------|----------------------|----------------------|----------------------|
| | (1) | (2) | (3) | <1/3 | 1/3<2/3 | >2/3 |
| <i>Origin country</i> | | | | | | |
| Individualism index | -0.083*** (0.017) | -0.070** (0.029) | -0.070* (0.036) | -0.142*** (0.050) | -0.098* (0.052) | -0.038 (0.053) |
| Gini coefficient | | -0.092 (0.061) | -0.092 (0.075) | -0.345*** (0.112) | -0.228* (0.130) | -0.130 (0.126) |
| Log GDP per capita | | 0.619 (0.818) | -0.275 (0.976) | 0.929 (1.459) | 2.039 (1.586) | -0.385 (1.519) |
| Mean social trust | | | -0.010 (0.047) | | | |
| Fractionalization | | | -0.079* (0.044) | | | |
| Polity score | | | 0.027 (0.028) | | | |
| <i>Residence country</i> | | | | | | |
| Individualism index | -0.203*** (0.031) | -0.138*** (0.037) | -0.123*** (0.043) | -0.113 (0.073) | -0.180** (0.083) | -0.267*** (0.061) |
| Gini coefficient | | 0.696*** (0.121) | 0.627*** (0.119) | 0.589* (0.328) | 0.694*** (0.224) | 0.318** (0.156) |
| Log GDP per capita | | -2.746 (1.962) | -3.070 (2.134) | -6.501* (3.611) | 1.278 (4.492) | 3.347 (3.648) |
| Mean social trust | | | -0.001 (0.039) | | | |
| Fractionalization | | | 0.020 (0.030) | | | |
| Polity score | | | -0.054 (0.099) | | | |
| <i>Individual characteristics</i> | | | | | | |
| Trust value | | 1.664** (0.753) | 1.775** (0.782) | 2.477 (1.696) | -1.485 (1.091) | -0.400 (1.119) |
| Life satisfaction value | | -0.064*** (0.015) | -0.069*** (0.016) | -0.010 (0.039) | -0.049** (0.024) | -0.043 (0.038) |
| Political left-right scale | | -0.175*** (0.019) | -0.171*** (0.020) | -0.209*** (0.036) | -0.176*** (0.032) | -0.137*** (0.026) |
| Education level | | -0.085*** (0.015) | -0.082*** (0.015) | -0.010 (0.027) | -0.081*** (0.023) | 0.002 (0.024) |
| Employment status | | -1.353*** (0.446) | -1.533*** (0.443) | -1.366 (1.057) | 0.594 (1.160) | -2.800** (1.077) |
| Household income | | -0.001** (0.000) | -0.001** (0.000) | -0.001*** (0.000) | -0.001*** (0.000) | -0.001* (0.000) |
| Time in new country | | 0.872** (0.336) | 0.864** (0.327) | | | |
| Sex | 2.669*** (0.218) | 1.637*** (0.612) | 1.727*** (0.630) | 2.623 (1.582) | 1.877 (1.254) | 1.986* (1.024) |
| Age | 0.125*** (0.013) | 0.061** (0.025) | 0.054** (0.025) | 0.008 (0.071) | 0.079 (0.052) | 0.071* (0.039) |
| Residence country FE | No | No | No | No | No | No |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Number of obs. | 52,767 | 8,759 | 8,470 | 1,211 | 1,503 | 1,960 |
| R-squared | 0.068 | 0.136 | 0.139 | 0.188 | 0.196 | 0.281 |

Notes: Robust standard errors clustered on origin country in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: Alesina et al. (2003); ESS (2016); EVS (2016); Hofstede et al. (2010); Marshall et al. (2016); Milanovic (2016); World Bank (2016); WVS (2016).

Figure 3: Estimated impact of individualism on preferences for redistribution by immigrant sample.



Note: Estimated standardized coefficients for individualism on preferences for redistribution, with 95 percent confidence intervals, controlling for origin country characteristics, individual characteristics, residence country and year fixed effects. *Source:* See Table A5 in the Appendix.

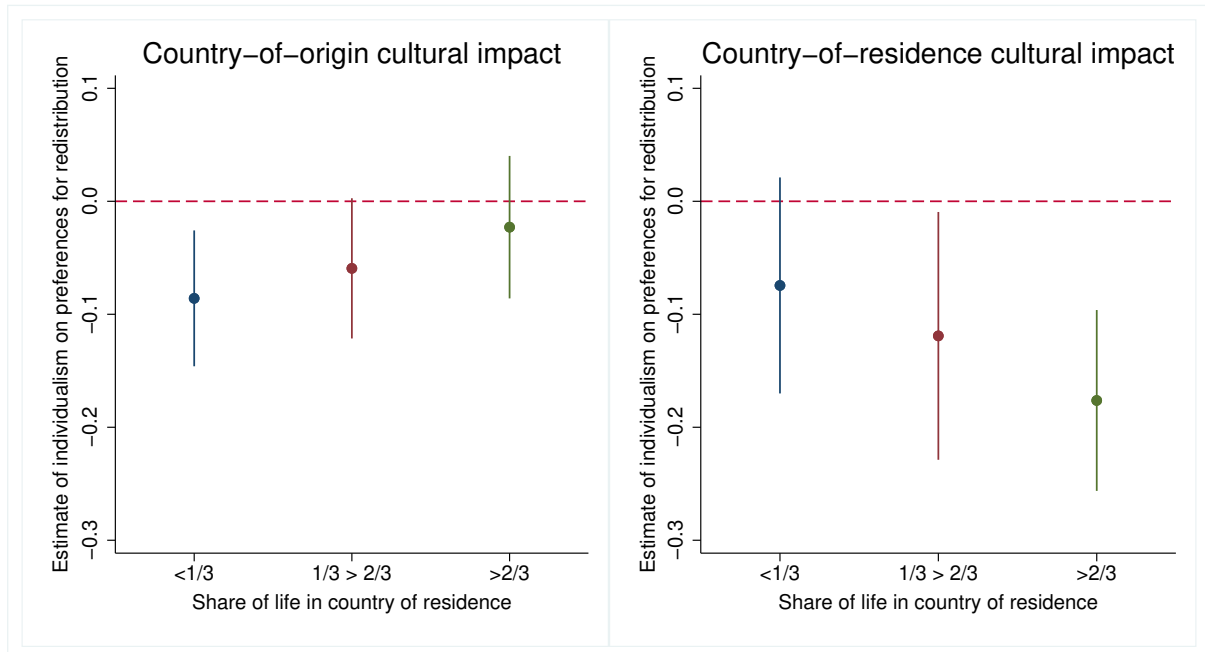
are presented in Figure 5, and indicate that there seems to be no statistically significant impact of the culture of origin on current preferences for redistribution if migration took place approximately before the age of 10. Culture in the current country of residence, however, seems to have a statistically significant impact regardless of the age at migration.

5 Robustness Checks

5.1 Pronoun-drop IV approach

Although none of the included country-of-origin controls are significant in the OLS regressions above, it could still be the case that there are some other omitted or unobserved variables driving the results. As an alternative strategy, I will thus use the pronoun-drop dummy as an instrument for individualism versus collectivism, i.e., to check whether it is actually individualism, rather than some other cultural variable in the country of origin,

Figure 4: Comparing country-of-origin versus country-of-residence impact of individualism on preferences for redistribution.



Note: Estimated standardized coefficients for individualism on preferences for redistribution, with 95 percent confidence intervals, controlling for origin country characteristics, residence country characteristics, individual characteristics and year fixed effects.

Source: See Table 7.

that drives the results. The assumption here is that the grammatical rule on pronoun drop affects preferences for redistribution only through its language-culture relationship, i.e., through its association with individualistic-collectivistic cultural beliefs. The results from these two-stage least squares (2SLS) regressions are shown in Table 5, where I have used the individual responses to the "language at home" question in the WVS (2016), the EVS (2016) and the ESS (2016), combined with pronoun-drop information from Abdurazokzoda & Davis (2016)'s new linguistic dataset.

The 2SLS regression results in Table 5 show that i) the pronoun drop seems to be a valid instrument for individualism-collectivism in the respect that it is negatively and statistically significantly related to the individualism index, and ii) using this instrument confirms the baseline results of a statistically significant negative association between individualism and an individual's preferences for redistribution. Here, the estimated impact is even stronger than for the OLS results, indicating that an increase of 10 percentage points on the individualism index is associated with a decrease of approximately 3 per-

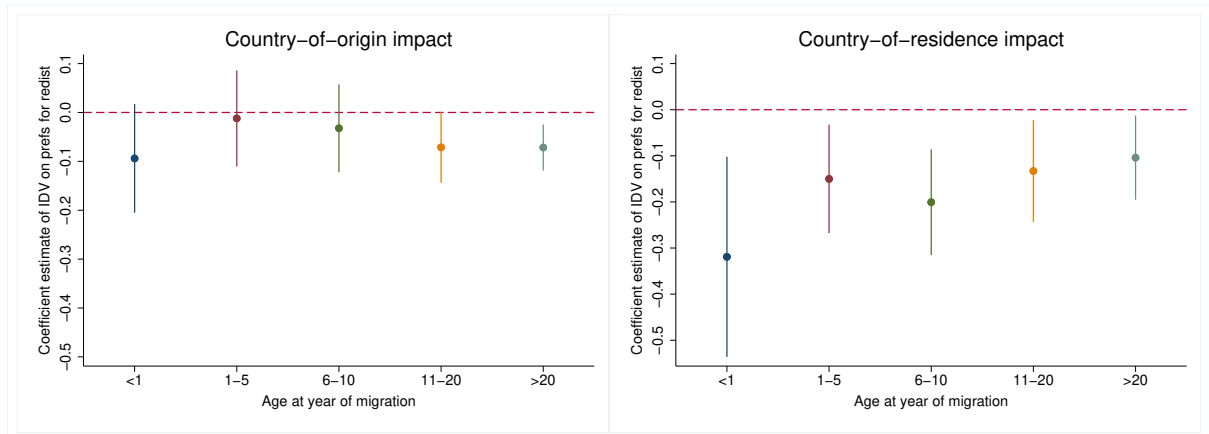
Table 5: IV regression results, immigrant sample.

| | Individualism (origin) | Preferences for redistribution | Individualism (origin) | Preferences for redistribution |
|-----------------------------------|---------------------------|-----------------------------------|---------------------------|-----------------------------------|
| | (1) <i>1st stage</i> | (1) <i>2nd stage</i> | (2) <i>1st stage</i> | (2) <i>2nd stage</i> |
| Individualism (origin) | | -0.355*** (0.065) | | -0.279*** (0.098) |
| Pronoun drop dummy | -12.168*** (4.085) | | -9.483*** (3.100) | |
| <i>Residence country</i> | | | | |
| Gini coefficient | | | -0.593*** (0.214) | 0.656*** (0.146) |
| Log GDP per capita | | | 12.188*** (2.943) | -6.457*** (2.134) |
| <i>Individual characteristics</i> | | | | |
| Trust value | | | 0.515 (0.664) | -0.889 (0.545) |
| Life satisfaction value | | | 0.030* (0.016) | -0.036*** (0.011) |
| Political left–right scale | | | -0.002 (0.011) | -0.179*** (0.023) |
| Education level | | | 0.014 (0.024) | -0.044*** (0.010) |
| Employment status | | | -0.483 (0.482) | -0.450 (0.443) |
| Household income | | | 0.000*** (0.000) | -0.001*** (0.000) |
| Sex | | | -0.210 (0.438) | 1.618*** (0.569) |
| Age | | | 0.106*** (0.034) | 0.129*** (0.022) |
| Number of obs. | 45,129 | 45,129 | 13,130 | 13,130 |
| R-squared | 0.079 | | 0.195 | |

Notes: Robust standard errors clustered on origin country in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Individualism index (IDV) measured in country of origin, and pronoun drop dummy for home language.

Sources: Abdurazokzoda & Davis (2016); ESS (2016); EVS (2016); Hofstede et al. (2010); Milanovic (2016); World Bank (2016); WVS (2016).

Figure 5: Comparing country-of-origin versus country-of-residence impact of individualism on preferences for redistribution by age at migration.



Note: Estimated standardized coefficients for individualism on preferences for redistribution, with 95 percent confidence intervals, controlling for origin country characteristics, residence country characteristics, individual characteristics and year fixed effects.

centage points on the preferences for redistribution scale. Moreover, since the baseline results are confirmed also when using this alternative individualism-collectivism measure, potential measurement error in the main survey-based individualism index does not seem to be driving the results.

5.2 Heterogeneity analyses

Using the epidemiological approach, another robustness check includes analyzing the different surveys separately (see Table A4 in the Appendix). The results hold for both surveys separately, but the individualism coefficient size is somewhat larger for the integrated EVS (2016) and WVS (2016) sample than for the ESS (2016). Since the wording of the survey question differs between these two samples, this indicates that the association between individualism and preferences for equality is slightly stronger than the association with preferences for redistribution.

Furthermore, I can also analyze the impact within different immigrant subsamples. These results are presented in Table 6.

The results in Table 6 show that the negative relation between individualism and preferences for income redistribution is robust to looking separately on immigrants who have

Table 6: Heterogeneity analysis, different immigrant samples.

| Preferences for redistribution | <i>More individualist origin</i> | <i>More collectivist origin</i> | <i>Between-region migration</i> | <i>Within-region migration</i> | <i>1st-generation immigrant</i> | <i>2nd-generation immigrant</i> |
|---------------------------------------|----------------------------------|---------------------------------|---------------------------------|--------------------------------|---------------------------------|---------------------------------|
| <i>Origin country</i> | | | | | | |
| Individualism index | -0.180*** (0.046) | -0.069** (0.034) | -0.078*** (0.029) | -0.065* (0.038) | -0.065** (0.026) | -0.035 (0.032) |
| Gini coefficient | -0.139 (0.148) | -0.047 (0.058) | 0.009 (0.067) | -0.042 (0.116) | -0.067 (0.054) | 0.016 (0.084) |
| Log GDP per capita | 6.255*** (1.551) | 0.395 (0.547) | -0.803 (0.737) | 1.451 (1.663) | 0.510 (0.572) | 0.321 (0.698) |
| <i>Individual characteristics</i> | | | | | | |
| Trust value | 0.360 (0.837) | 1.453 (0.995) | 2.179** (0.973) | 0.727 (1.083) | 1.085 (0.786) | -0.092 (0.588) |
| Life satisfaction value | -0.077** (0.031) | -0.047*** (0.014) | 0.008 (0.023) | -0.085*** (0.020) | -0.058*** (0.013) | -0.084*** (0.014) |
| Political left–right scale | -0.206*** (0.026) | -0.165*** (0.025) | -0.146*** (0.023) | -0.196*** (0.022) | -0.179*** (0.019) | -0.187*** (0.025) |
| Education level | -0.086*** (0.021) | -0.075*** (0.015) | -0.045*** (0.014) | -0.092*** (0.016) | -0.086*** (0.011) | -0.085*** (0.012) |
| Employment status | -0.320 (0.625) | -0.783 (0.483) | -1.011 (0.674) | -0.512 (0.399) | -0.787* (0.439) | -0.570 (0.477) |
| Household income | -0.001** (0.000) | -0.001*** (0.000) | -0.001*** (0.000) | -0.001* (0.000) | -0.001*** (0.000) | -0.001*** (0.000) |
| Time in new country | 0.629 (0.492) | 1.362*** (0.445) | 2.166*** (0.511) | 0.898** (0.407) | | |
| Sex | 2.771*** (0.951) | 1.717*** (0.549) | 1.405 (0.881) | 2.004*** (0.557) | 1.930*** (0.411) | 2.730*** (0.526) |
| Age | 0.073* (0.038) | 0.054** (0.023) | 0.010 (0.027) | 0.065** (0.024) | 0.097*** (0.015) | 0.100*** (0.018) |
| Residence country FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Number of obs. | 3,349 | 8,736 | 3,919 | 8,224 | 13,161 | 10,705 |
| R-squared | 0.206 | 0.139 | 0.156 | 0.157 | 0.160 | 0.182 |

Notes: Robust standard errors clustered on origin country in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); Milanovic (2016); World Bank (2016); WVS (2016).

emigrated from a more individualistic culture relative to their country of destination, as well as from a relatively more collectivistic culture, respectively; from another country within the same geographical region, and those who have emigrated from another geographical region, respectively.⁷ The relationship, however, is strongest among those who have migrated from one region to another, and to a relatively more collectivist region. The association is also robust to looking only on the sample of first-generation immigrants (i.e., those with another nationality and/or country of birth than their country of residence). Analyzing the sample of second-generation immigrants separately, however, there seems to be no statistically significant impact of individualism in the parents' country of origin on their children's preferences for redistribution. These results again indicate the existence of an assimilation or integration process such that the cultural impact of individualism on redistributive preferences is not persistent across generations and possibly weakens off with time spent in the new institutional and cultural environment. Notably, this insignificant relationship among the second-generation immigrants is in contrast to what has been found by, e.g., [Luttmer & Singhal \(2011\)](#).

5.3 Matching estimators

As a final sensitivity analysis, I will also check if the results are robust to using matching as an alternative estimation strategy. I hence use the propensity-score matching method, where I compare individuals that are similar in a number of observable characteristics but differ in their individualism versus collectivism cultural belonging. I thus create a dummy variable in which I define individuals that have a country-of-origin individualism index value above 50 as individualists and those with a value below 50 as collectivists. For a comparison of the mean values in the two samples, see [Table A5](#) in the Appendix. The estimation results when matching and comparing immigrant individualists to collectivists are shown in [Table 7](#).

As seen in [Table 7](#), using this matching method I also find a negative and statistically significant relationship between individualism and preferences for redistribution. More

⁷Using the United Nations' classification of regions.

Table 7: Propensity-score matching results, immigrant sample.

| Preferences for redistribution | <i>ATE</i> |
|---------------------------------------|---------------------|
| Individualist dummy | -1.267** (0.569) |
| Number of obs. | 26,743 |

Notes: Average treatment effects (ATE). Robust standard errors in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Matching variables are individual-level trust value, life satisfaction value, political left-right scale, education level, employment status, household income, sex, age, and residence country FE.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); Milanovic (2016); World Bank (2016); WVS (2016).

specifically, the average treatment effect of having an individualistic, as opposed to collectivistic, culture is a reduction of approximately 1.3 percentage points on the preferences for redistribution scale.

6 Conclusion

In this study, I have analyzed the association between individualism versus collectivism and individuals' preferences for income redistribution and equality, using variation in immigrants' country of origin to separate the effect of culture from the otherwise endogenously determined institutional environment. Doing so, I have found strong support for a negative relationship between individualistic cultural beliefs and redistributive preferences, i.e., individualistic societies seem to foster preferences for income inequality. These results were confirmed using matching estimators as well as the grammatical rule on pronoun drop as a linguistic instrument for individualism-collectivism. Heterogeneity analyses also showed that the cultural impact of individualism on redistributive preferences is not significantly persistent over generations. Moreover, the impact of cultural origin only seems to be statistically significant if migration took place after the age of 10. More research would be needed in order to better understand the workings of such cultural adaption and its relation to institutional and cultural change. It would also be interesting to analyze the association between other cultural dimensions and egalitarian preferences, as well as the impact of individualism on actual redistribution and income equality.

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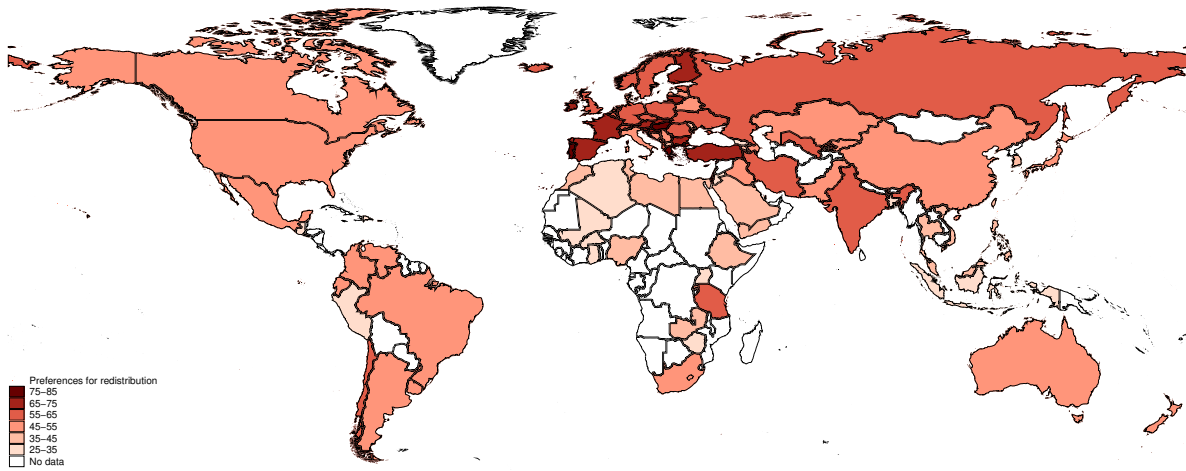
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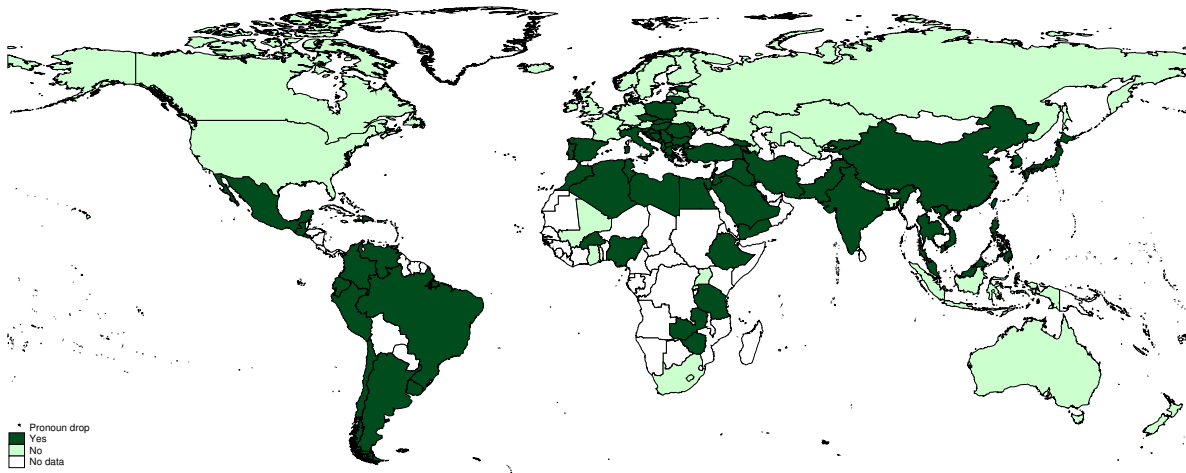
Appendix

Figure A1: Mean preferences for redistribution around the world.



Sources: ESS (2016); EVS (2016); WVS (2016).

Figure A2: Mean pronoun drop around the world.



Source: Abdurazokzoda & Davis (2016).

Table A1: Correlation matrix.

| | Pref's for redist. | IDV (origin) | Mean social trust (origin) | Mean pref's for redist. (origin) |
|-----------------------------------|-----------------------|-----------------|-------------------------------|-------------------------------------|
| <i>Origin country</i> | | | | |
| Individualism index | -0.080*** | | | |
| Mean social trust | -0.081*** | 0.608*** | | |
| Mean pref's for redist. | 0.002 | 0.119*** | 0.225*** | |
| Gini coefficient | 0.015*** | -0.237*** | -0.353*** | -0.227*** |
| Log GDP per capita | -0.068*** | 0.706*** | 0.600*** | 0.475*** |
| Fractionalization | -0.018** | -0.234*** | -0.112*** | -0.347*** |
| Polity score | -0.091*** | 0.612*** | 0.496*** | 0.414*** |
| <i>Individual characteristics</i> | | | | |
| Trust value | 0.054*** | 0.134*** | 0.196*** | -0.020*** |
| Life satisfaction value | -0.051*** | 0.148*** | 0.138*** | 0.004 |
| Political left–right scale | -0.131*** | -0.016*** | 0.007 | -0.037*** |
| Education level | -0.143*** | 0.057*** | 0.106*** | -0.018*** |
| Employment status | -0.069*** | 0.029*** | 0.055*** | 0.026*** |
| Household income | -0.089*** | 0.168*** | 0.188*** | 0.041*** |
| Time in new country | 0.077*** | 0.031*** | 0.068*** | 0.017*** |
| Sex | 0.050*** | -0.011*** | 0.003 | -0.005 |
| Age | 0.122*** | 0.055*** | 0.035*** | 0.019*** |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: Alesina et al. (2003); ESS (2016); EVS (2016); Hofstede et al. (2010); Marshall et al. (2016); Milanovic (2016); World Bank (2016); WVS (2016).

Table A2: OLS, ordered logistic and ordered probit regression results, full sample.

| Preferences for redistribution | <i>OLS</i> | <i>Ordered logit</i> | <i>Ordered probit</i> |
|---------------------------------------|----------------------|----------------------|-----------------------|
| <i>Residence country</i> | | | |
| Individualism index | -0.144*** (0.006) | -0.010*** (0.000) | -0.006*** (0.000) |
| Gini coefficient | 0.262*** (0.018) | 0.017*** (0.001) | 0.010*** (0.001) |
| Log GDP per capita | 0.362 (0.229) | -0.006 (0.015) | 0.006 (0.009) |
| <i>Individual characteristics</i> | | | |
| Trust value | -0.395** (0.156) | -0.045*** (0.010) | -0.024*** (0.006) |
| Life satisfaction value | -0.089*** (0.004) | -0.007*** (0.000) | -0.004*** (0.000) |
| Political left–right scale | -0.181*** (0.003) | -0.012*** (0.000) | -0.007*** (0.000) |
| Education level | -0.111*** (0.003) | -0.007*** (0.000) | -0.004*** (0.000) |
| Employment status | -1.148*** (0.130) | -0.075*** (0.009) | -0.044*** (0.005) |
| Household income | -0.001*** (0.000) | 0.000*** (0.000) | 0.000*** (0.000) |
| Sex | 2.656*** (0.144) | 0.154*** (0.009) | 0.092*** (0.006) |
| Age | 0.055*** (0.004) | 0.004*** (0.000) | 0.002*** (0.000) |
| Residence country FE | No | No | No |
| Year FE | Yes | Yes | Yes |
| Number of obs. | 147,110 | 147,110 | 147,110 |
| R-squared | 0.146 | 0.037 | 0.036 |

Notes: Robust standard errors in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); Milanovic (2016); World Bank (2016); WVS (2016).

Table A3: Sensitivity analysis, different immigrant samples.

| Preferences for redistribution | <i>Citizenship</i> | <i>Birth country</i> | <i>Mother's origin</i> | <i>Father's origin</i> |
|---------------------------------------|----------------------|----------------------|------------------------|------------------------|
| <i>Origin country</i> | | | | |
| Individualism index | -0.098** (0.043) | -0.065** (0.027) | -0.035 (0.023) | -0.026 (0.024) |
| Gini coefficient | -0.165** (0.082) | -0.062 (0.060) | -0.040 (0.056) | -0.048 (0.058) |
| Log GDP per capita | 0.597 (1.232) | 0.574 (0.639) | 0.535 (0.573) | 0.229 (0.583) |
| <i>Individual characteristics</i> | | | | |
| Trust value | 0.499 (1.016) | 1.142 (0.752) | 0.476 (0.615) | 0.163 (0.750) |
| Life satisfaction value | -0.062*** (0.022) | -0.055*** (0.012) | -0.066*** (0.016) | -0.070*** (0.014) |
| Political left–right scale | -0.201*** (0.025) | -0.176*** (0.019) | -0.168*** (0.020) | -0.170*** (0.021) |
| Education level | -0.091*** (0.013) | -0.084*** (0.012) | -0.085*** (0.012) | -0.081*** (0.013) |
| Employment status | -0.096 (0.516) | -0.699* (0.396) | -0.466 (0.372) | -0.597 (0.420) |
| Household income | -0.001*** (0.000) | -0.001*** (0.000) | -0.001*** (0.000) | -0.001*** (0.000) |
| Sex | 1.512** (0.719) | 1.976*** (0.482) | 2.117*** (0.383) | 2.568*** (0.408) |
| Age | 0.107*** (0.019) | 0.087*** (0.018) | 0.080*** (0.015) | 0.081*** (0.014) |
| Residence country FE | Yes | Yes | Yes | Yes |
| Year FE | Yes | Yes | Yes | Yes |
| Number of obs. | 5,662 | 12,110 | 15,976 | 16,360 |
| R-squared | 0.173 | 0.153 | 0.168 | 0.163 |

Notes: Robust standard errors clustered on origin country in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); Milanovic (2016); World Bank (2016); WVS (2016).

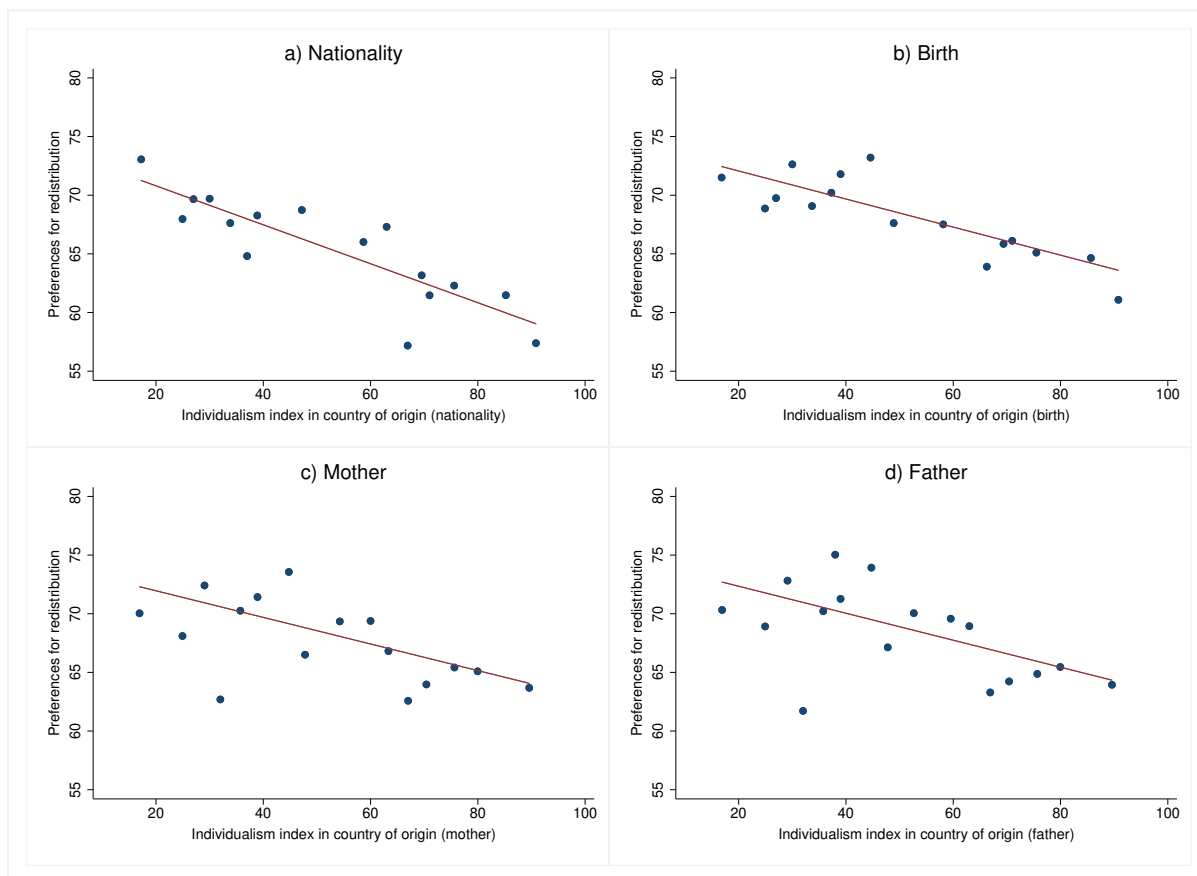
Table A4: Robustness analysis, different surveys.

| Preferences for redistribution | <i>Incomes more equal (EVS & WVS)</i> | <i>Government redistribution (ESS)</i> |
|---------------------------------------|---|--|
| <i>Origin country</i> | | |
| Individualism index | -0.110** (0.046) | -0.059** (0.029) |
| Gini coefficient | -0.305*** (0.094) | -0.038 (0.065) |
| Log GDP per capita | 0.165 (1.441) | 0.669 (0.677) |
| <i>Individual characteristics</i> | | |
| Trust value | 1.350 (1.148) | -0.434 (0.673) |
| Life satisfaction value | -0.023 (0.031) | -0.033*** (0.011) |
| Political left–right scale | -0.136*** (0.034) | -0.191*** (0.020) |
| Education level | -0.094*** (0.028) | -0.049*** (0.010) |
| Employment status | -0.573 (0.991) | -0.388 (0.395) |
| Household income | -0.001 (0.001) | -0.001*** (0.000) |
| Time in new country | 0.351 (1.267) | 1.566*** (0.320) |
| Sex | 3.131** (1.189) | 1.736*** (0.545) |
| Age | 0.057 (0.035) | 0.045** (0.020) |
| Residence country FE | Yes | Yes |
| Year FE | Yes | Yes |
| Number of obs. | 2,355 | 9,788 |
| R-squared | 0.145 | 0.139 |

Notes: Robust standard errors clustered on origin country in parentheses; * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); Milanovic (2016); World Bank (2016); WVS (2016).

Figure A3: Correlation between individualism and preferences for redistribution by cultural origin source: binned scatterplots.



Note: Bins based on a) 13,374, b) 28,806, c) 37,119 and d) 37,963 observations in total.

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010); WVS (2016).

Table A5: Means comparison, immigrant sample.

| | Individualists (IDV > 50) | | | Collectivists (IDV < 50) | | |
|--------------------------------|---------------------------|-------|-----------|--------------------------|-------|-----------|
| | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. |
| Preferences for redistribution | 29,645 | 66.64 | 28.47 | 31,841 | 70.99 | 28.32 |
| Individualism index (origin) | 25,714 | 68.98 | 10.23 | 30,925 | 34.23 | 9.10 |
| Trust value | 24,589 | 0.54 | 0.50 | 26,303 | 0.43 | 0.50 |
| Life satisfaction value | 30,202 | 71.16 | 22.62 | 32,617 | 64.10 | 25.19 |
| Political left–right scale | 26,645 | 49.31 | 22.44 | 25,264 | 50.35 | 23.61 |
| Education level | 25,071 | 52.19 | 30.63 | 27,762 | 50.73 | 30.48 |
| Employment status | 30,131 | 1.44 | 0.62 | 32,690 | 1.42 | 0.64 |
| Household income | 23,348 | 2,660 | 2,562 | 23,320 | 1,988 | 2,193 |
| Sex | 30,332 | 0.55 | 0.50 | 32,885 | 0.56 | 0.50 |
| Age | 30,211 | 46.78 | 17.89 | 32,739 | 45.76 | 17.85 |

Sources: ESS (2016); EVS (2016); Hofstede et al. (2010).